

# Smooth Calibration, Leaky Forecasts, Finite Recall, and Nash Dynamics\*

Dean P. Foster<sup>†</sup>      Sergiu Hart<sup>‡</sup>

December 1, 2017

## Abstract

We propose to smooth out the calibration score, which measures how good a forecaster is, by combining nearby forecasts. While regular calibration can be guaranteed only by randomized forecasting procedures, we show that *smooth calibration* can be guaranteed by *deterministic* procedures. As a consequence, it does not matter if the forecasts are *leaked*, i.e., made known in advance: smooth calibration can nevertheless be guaranteed (while regular calibration cannot). Moreover, our procedure has finite recall, is stationary, and all forecasts lie on a finite grid. To construct the procedure, we deal also with the related setups of online linear regression and weak calibration. Finally, we show that smooth calibration yields uncoupled finite-memory *dynamics* in  $n$ -person games—“smooth calibrated learning”—in which the players play approximate *Nash equilibria* in almost all periods (by contrast, calibrated learning, which uses regular calibration, yields

---

\*Previous versions: July 2012, February 2015, March 2017. Research of the second author was partially supported by a European Research Council (ERC) Advanced Investigator grant. The authors thank Yakov Babichenko for useful comments, and the editor, associate editor, and referees for their very careful reading and helpful suggestions.

<sup>†</sup>Amazon Inc, New York City, and University of Pennsylvania, Philadelphia. *e-mail*: [dean@foster.net](mailto:dean@foster.net) *web page*: <http://deanfoster.net/>

<sup>‡</sup>Institute of Mathematics, Department of Economics, and Center for the Study of Rationality, The Hebrew University of Jerusalem. *e-mail*: [hart@huji.ac.il](mailto:hart@huji.ac.il) *web page*: <http://www.ma.huji.ac.il/hart>

only that the time-averages of play are approximate correlated equilibria).

## Contents

|          |   |           |
|----------|---|-----------|
| <b>1</b> | <b>Introduction</b>                             | <b>2</b>  |
| 1.1      | Literature . . . . .                            | 5         |
| <b>2</b> | <b>Calibration: Model and Result</b>            | <b>6</b>  |
| 2.1      | The Calibration Game . . . . .                  | 6         |
| 2.2      | Smooth Calibration . . . . .                    | 7         |
| 2.3      | Leaky Forecasts . . . . .                       | 9         |
| 2.4      | Result . . . . .                                | 11        |
| <b>3</b> | <b>Online Linear Regression</b>                 | <b>11</b> |
| 3.1      | Forward Algorithm . . . . .                     | 13        |
| 3.2      | Discounted Forward Algorithm . . . . .          | 14        |
| 3.3      | Windowed Discounted Forward Algorithm . . . . . | 18        |
| <b>4</b> | <b>Weak Calibration</b>                         | <b>23</b> |
| <b>5</b> | <b>Smooth Calibration</b>                       | <b>29</b> |
| <b>6</b> | <b>Nash Equilibrium Dynamics</b>                | <b>34</b> |
|          | <b>References</b>                               | <b>43</b> |
| <b>A</b> | <b>Appendix</b>                                 | <b>46</b> |

## 1 Introduction

How good is a forecaster? Assume for concreteness that every day the forecaster issues a forecast of the type “the chance of rain tomorrow is 30%.” A simple test one may conduct is to calculate the proportion of rainy days out of those days for which the forecast was 30%, and compare it to 30%; and do

the same for all other forecasts. A forecaster is said to be *calibrated* if, in the long run, the differences between the actual proportions of rainy days and the forecasts are small—no matter what the weather really was (see Dawid 1982).

What if rain is replaced by an event that is under the control of another agent? If the forecasts are made public before the agent decides on his action—we refer to this setup as “*leaky forecasts*”—then calibration *cannot* be guaranteed; for example, the agent can make the event happen if and only if the forecast is less than 50%, and so the forecasting error (that is, the “calibration score”) is always at least 50%. However, if in each period the forecast and the agent’s decision are made “simultaneously”—which means that neither one knows the other’s decision before making his own—then calibration *can* be guaranteed; see Foster and Vohra (1998). The procedure that yields calibration no matter what the agent’s decisions are requires the use of *randomizations* (e.g., with probability 1/2 the forecaster announces 30%, and with probability 1/2 he announces 60%). Indeed, as the discussion at the beginning of this paragraph suggests, one cannot have a deterministic procedure that is calibrated (see Dawid 1985 and Oakes 1985).

Now the standard calibration score is overly fastidious: the days when the forecast was, say, 30.01% are considered separately from the days when the forecast was 29.99% (formally, the calibration score is a highly discontinuous function of the data, i.e., the forecasts and the actions). This suggests that one first combines all days when the forecast was *close to* 30%, and only then compares the 30% with the proportion of rainy days. If, say, there were 200 days with a forecast of 30.01%, out of which 10 were rainy, and another 100 days with a forecast of 29.99%, out of which 80 were rainy, then the forecaster is very far from being calibrated; however, he is smoothly calibrated, as his forecasts were all close to 30%, and there were  $90/300 = 30\%$  rainy days. Undershooting at 29.99% and overshooting at 30.01% is now balanced out. Formally, what this amounts to is applying a so-called “smoothing” operation to the forecasting errors (which makes smooth calibration easier to obtain than calibration).<sup>1</sup>

---

<sup>1</sup>Corollary 12 in Section 4 will formally show that regular calibration implies smooth

Perhaps surprisingly, once we consider smooth calibration, there is no longer a need for randomization when making the forecasts: we will show that there exist *deterministic* procedures that guarantee smooth calibration, no matter what the agent does. In particular, it follows that it does not matter if the forecasts are made known to the agent before his decision, and so smooth calibration can be guaranteed even when forecasts may be leaked.<sup>2</sup> This may come as a surprise, because, as pointed out above, an agent who knows the forecast *before* deciding on the weather will choose rain when the forecast is less than 50% and no rain otherwise, giving a calibration error of 50% or more, no matter what the forecaster does. However, against such an agent one can easily be *smoothly* calibrated, by forecasting 50.01% on odd days and 49.99% on even days (the resulting weather will then alternate between rain and no rain, and so half the days will be rainy days—and all the forecasts are indeed close to 50%). What this proves is only that one can be smoothly calibrated against this specific strategy of the agent (this is the strategy that shows that it is impossible to have calibration with deterministic leaky procedures); our result shows that one can in fact *guarantee* smooth calibration with a deterministic strategy, against *any* strategy of the agent.

The forecasting procedure that we construct and that guarantees smooth calibration has moreover finite recall (i.e., only the forecasts and actions of the last  $R$  periods are taken into account, for some fixed finite  $R$ ), and is stationary (i.e., independent of “calendar time”: the forecast is the same any time that the “window” of the past  $R$  periods is the same).<sup>3</sup> Finally, we can have all the forecasts lie on some finite fixed grid.

The construction starts with the “online linear regression” problem, introduced by Foster (1991), where one wants to generate every period a good linear estimator based only on the data up to that point. We provide a finite-recall stationary algorithm for this problem; see Section 3. We then

---

calibration.

<sup>2</sup>When the forecasting procedure is deterministic it can be simulated by the agent, and so it is irrelevant whether the agent *observes* the forecasts, or just *computes* them by himself, before taking his action.

<sup>3</sup>Another, seemingly less elegant, way to obtain this is by restarting the procedure once in a while; see, e.g., Lehrer and Solan (2009).

use this algorithm, together with a fixed-point argument, to obtain “weak calibration,” a concept introduced by Kakade and Foster (2004) and Foster and Kakade (2006); see Section 4. Section 5 shows that weak and smooth calibration are essentially equivalent, which yields the existence of smoothly calibrated procedures. Finally, these procedures are used to obtain dynamics (“smoothly calibrated learning”) that are uncoupled, have finite memory, and are close to Nash equilibria most of the time (while the similar dynamics that are based on regular calibration yield only the time average becoming close to correlated equilibria; see Foster and Vohra 1997).

## 1.1 Literature

The *calibration problem* has been extensively studied, starting with Dawid (1982), Oakes (1985), and Foster and Vohra (1998); see Olszewski (2015) for a comprehensive survey of the literature. Kakade and Foster (2004) and Foster and Kakade (2006) introduced the notion of *weak calibration*, which shares many properties with smooth calibration. In particular, both can be guaranteed by deterministic procedures, and both are of the “general fixed point” variety: they can find fixed points of arbitrary continuous functions (see for instance the last paragraph in Section 2.3).<sup>4</sup> However, while weak calibration may be at times technically more convenient to work with, smooth calibration is the more natural concept, easier to interpret and understand; it is, after all, just a standard smoothing of regular calibration.

The *online regression problem*—see Section 3 for details—was introduced by Foster (1991); for further improvements, see J. Foster (1999), Vovk (2001), Azoury and Warmuth (2001), and the book of Cesa-Bianchi and Lugosi (2006).

---

<sup>4</sup>They are thus more “powerful” than the standard calibration procedures (such as those based on Blackwell’s approachability), which find *linear* fixed points (such as eigenvectors and invariant probabilities).

## 2 Calibration: Model and Result

In this section we present the calibration game in its standard and “leaky” versions, introduce the notion of smooth calibration, and state our main results.

### 2.1 The Calibration Game

Let<sup>5</sup>  $C \subseteq \mathbb{R}^m$  be a compact convex set, and let  $A \subseteq C$  (for example,  $C$  could be the set of probability distributions  $\Delta(A)$  over a finite set  $A$ , which is identified with the set of unit vectors in  $C$ , or a product of such sets). The *calibration game* has two players: the “action” player—the “A-player” for short—and the “conjecture” (or “calibrating”) player—the “C-player” for short. At each time period  $t = 1, 2, \dots$ , the C-player chooses  $c_t \in C$  and the A-player chooses  $a_t \in A$ . There is full monitoring and perfect recall: at time  $t$  both players know the realized history  $h_{t-1} = (c_1, a_1, \dots, c_{t-1}, a_{t-1}) \in (C \times A)^{t-1}$ .

In the *standard* calibration game,  $c_t$  and  $a_t$  are chosen simultaneously (perhaps in a mixed, i.e., randomized, way). In the *leaky* calibration game,  $a_t$  is chosen after  $c_t$  has been chosen and revealed; thus,  $c_t$  is a function of  $h_{t-1}$ , whereas  $a_t$  is a function of  $h_{t-1}$  and  $c_t$ . Formally, a pure strategy of the C-player is  $\sigma : \cup_{t \geq 1} (C \times A)^{t-1} \rightarrow C$ , and a pure strategy of the A-player is  $\tau : \cup_{t \geq 1} (C \times A)^{t-1} \rightarrow A$  in the standard game, and  $\tau : \cup_{t \geq 1} (C \times A)^{t-1} \times C \rightarrow A$  in the leaky game. A pure strategy of the C-player will also be referred to as *deterministic*.

The calibration score—which the C-player wants to minimize—is defined at time  $T \geq 1$  as follows. For every forecast  $c$  in  $C$  let  $n(c) \equiv n_T(c) := |\{1 \leq t \leq T : c_t = c\}|$  be the number of times that it has been used, and let

$$\bar{a}(c) \equiv \bar{a}_T(c) := \frac{1}{n(c)} \sum_{s=1}^T \mathbf{1}_{c_s=c} a_s$$

be the average of the actions in the periods when the forecast was  $c$ , where

---

<sup>5</sup>We denote by  $\mathbb{R}^m$  the  $m$ -dimensional Euclidean space, with the usual  $\ell_2$ -norm  $\|\cdot\|$ .

we write  $\mathbf{1}_{x=y}$  for the indicator that  $x = y$  (i.e.,  $\mathbf{1}_{x=y} = 1$  when  $x = y$  and  $\mathbf{1}_{x=y} = 0$  otherwise);  $\bar{a}(c)$  is defined only when  $c$  appears in the sequence  $c_1, \dots, c_T$ , i.e.,  $n(c) > 0$ . The *calibration score* at time  $T$  is then defined as<sup>6</sup>

$$K_T := \sum_{c \in C} \frac{n(c)}{T} \|\bar{a}(c) - c\|.$$

Thus  $K_T$  is the mean distance between the forecast  $c$  and the average  $\bar{a}(c)$  of the actions  $a$  in those periods where the forecast was  $c$ , weighted proportionately to how often each forecast  $c$  has been used in those  $T$  periods. An alternative formulation is easily seen to be

$$K_T = \frac{1}{T} \sum_{t=1}^T \|\bar{a}_t - c_t\|, \quad (1)$$

where  $\bar{a}_t := \bar{a}(c_t)$ , i.e.,

$$\bar{a}_t := \frac{\sum_{s=1}^T \mathbf{1}_{c_s=c_t} a_s}{\sum_{s=1}^T \mathbf{1}_{c_s=c_t}};$$

indeed, for each  $c$  there are  $n(c)$  identical terms in (1) that each equal  $\|\bar{a}(c) - c\|$ .

## 2.2 Smooth Calibration

We introduce the notion of “smooth calibration.” A *smoothing function* is a function  $\Lambda : C \times C \rightarrow [0, 1]$  with  $\Lambda(c, c) = 1$  for every  $c$ . Its interpretation is that  $\Lambda(c', c)$  gives the weight that we assign to  $c'$  when we are at  $c$ . We will use  $\Lambda(c', c)$  instead of the indicator  $\mathbf{1}_{c'=c}$  to “smooth” out the forecasts and the average actions. Specifically, put

$$\bar{a}_t^\Lambda := \frac{\sum_{s=1}^T \Lambda(c_s, c_t) a_s}{\sum_{s=1}^T \Lambda(c_s, c_t)} \quad \text{and} \quad c_t^\Lambda := \frac{\sum_{s=1}^T \Lambda(c_s, c_t) c_s}{\sum_{s=1}^T \Lambda(c_s, c_t)}.$$

The  $\Lambda$ -*smoothed calibration score* at time  $T$  is then defined as

---

<sup>6</sup>The summation is over all  $c$  with  $n(c) > 0$ , and we use the Euclidean norm (the squared distance  $\|\bar{a}(c) - c\|^2$  may well be used instead, in line with standard statistics usage).

$$K_T^\Lambda = \frac{1}{T} \sum_{t=1}^T \|\bar{a}_t^\Lambda - c_t^\Lambda\|. \quad (2)$$

A standard (and useful) assumption is a Lipschitz condition: there exists  $L < \infty$  such that  $|\Lambda(c', c) - \Lambda(c'', c)| \leq L\|c' - c''\|$  for all  $c, c', c'' \in C$ . Thus, the functions  $\Lambda(\cdot, c)$  are uniformly Lipschitz:  $\mathcal{L}(\Lambda(\cdot, c)) \leq L$  for every  $c \in C$ , where  $\mathcal{L}(f) := \sup\{\|f(x) - f(y)\| / \|x - y\| : x, y \in X, x \neq y\}$  denotes the *Lipschitz constant* of the function  $f$  (if  $f$  is not a Lipschitz function then  $\mathcal{L}(f) = +\infty$ ; when  $\mathcal{L}(f) \leq L$  we say that  $f$  is *L-Lipschitz*).

Two classic examples of Lipschitz smoothing functions are: (i) the so-called *tent function*  $\Lambda(c', c) = [1 - \|c' - c\|/\delta]_+$  for  $\delta > 0$ , where  $[z]_+ := \max\{z, 0\}$ ; thus, only points  $c'$  within distance  $\delta$  of  $c$  are considered, and their weight is proportional to the distance from  $c$  (see Figure 1 *Right* for

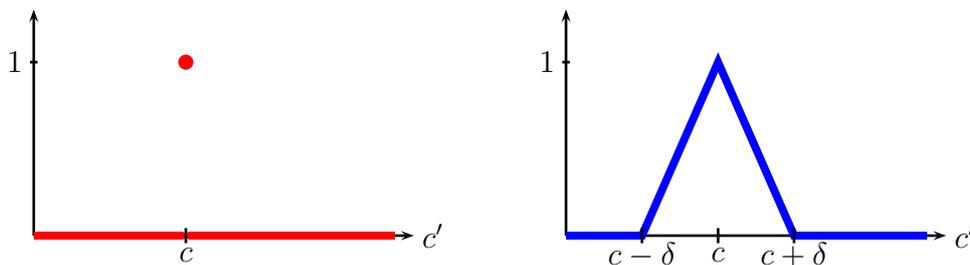


Figure 1: *Left*: The indicator function  $\mathbf{1}_{x=c}$ . *Right*: The  $\delta$ -tent smoothing function  $\Lambda(x, c) = [1 - \|x - c\|/\delta]_+$  with Lipschitz bound  $L = 1/\delta$ .

this function  $\Lambda$ , and compare it with the indicator function in Figure 1 *Left*); and (ii) the so-called *Gaussian* (or *normal*) smoothing function  $\Lambda(c', c) = \exp(-\|c' - c\|^2/(2\sigma^2))$ .

**Remarks.** (a) The original calibration score  $K_T$  is obtained when  $\Lambda$  is the indicator function, i.e.,  $\Lambda(c', c) = \mathbf{1}_{c'=c}$  for all  $c, c' \in C$ .

(b) The normalization  $\Lambda(c, c) = 1$  pins down the Lipschitz constant (otherwise one could replace  $\Lambda$  with  $\alpha\Lambda$  for small  $\alpha > 0$ , and so lower the Lipschitz constant without affecting the score).

(c) Smoothing both  $\bar{a}_t$  and  $c_t$  and then taking the difference is the same

as smoothing the difference:  $\bar{a}_t^\Lambda - c_t^\Lambda = (\bar{a}_t - c_t)^\Lambda$ . Moreover, smoothing  $a_t$  is the same as smoothing  $\bar{a}_t$ , i.e.,  $\bar{a}_t^\Lambda = a_t^\Lambda$ .

(d) An alternative score smoothes only the average action  $\bar{a}_t$ , but not the forecast  $c_t$ :

$$\tilde{K}_T^\Lambda = \frac{1}{T} \sum_{t=1}^T \|\bar{a}_t^\Lambda - c_t\|.$$

If the smoothing function puts positive weight only in small neighborhoods, i.e., there is  $\delta > 0$  such that  $\Lambda(c', c) > 0$  only when  $\|c' - c\| \leq \delta$ , then the difference between  $K_T^\Lambda$  and  $\tilde{K}_T^\Lambda$  is at most  $\delta$  (because in this case  $\|c_t^\Lambda - c_t\| \leq \delta$  for every  $t$ ). More generally,  $|K_T^\Lambda - \tilde{K}_T^\Lambda| \leq \delta$  when  $(1/T) \sum_{t=1}^T \|c_t^\Lambda - c_t\| \leq \delta$  for any collection of points  $c_1, \dots, c_T \in C$ , which is indeed the case, for instance, for the Gaussian smoothing with small enough  $\sigma^2$ . The reason that we prefer to use  $K^\Lambda$  rather than  $\tilde{K}^\Lambda$  is that  $K^\Lambda$  vanishes when there is perfect calibration (i.e.,  $\bar{a}_t = c_t$  for all  $t$ ), whereas  $\tilde{K}^\Lambda$  remains positive; clean statements such as  $K_T^\Lambda \leq \varepsilon$  become  $\tilde{K}_T^\Lambda \leq \varepsilon + \delta$ .

Finally, given  $\varepsilon > 0$  and  $L < \infty$ , we will say that a strategy of the C-player—which is also called a “procedure”—is  $(\varepsilon, L)$ -smoothly calibrated if there is  $T_0 \equiv T_0(\varepsilon, L)$  such that

$$K_T^\Lambda = \frac{1}{T} \sum_{t=1}^T \|\bar{a}_t^\Lambda - c_t^\Lambda\| \leq \varepsilon \tag{3}$$

holds almost surely, for every strategy of the A-player, every  $T > T_0$ , and every smoothing function  $\Lambda : C \times C \rightarrow [0, 1]$  that is  $L$ -Lipschitz in the first coordinate. Unlike standard calibration, which can be guaranteed only with high probability, smooth calibration may be obtained by deterministic procedures—as will be shown below—in which case we may well require (3) to always hold (rather than just almost surely).

## 2.3 Leaky Forecasts

We will say that a procedure (i.e., a strategy of the C-player) is *leaky* (smoothly) calibrated if it is (smoothly) calibrated also in the leaky setup, that is, against

an A-player who may choose his action  $a_t$  at time  $t$  depending on the forecast  $c_t$  made by the C-player at time  $t$  (i.e., the A-player moves after the C-player). While, as we saw in the Introduction, there are no leaky calibrated procedures, we will show that there are *leaky smoothly calibrated* procedures.

Deterministic procedures (i.e., pure strategies of the C-player) are clearly leaky: the A-player can use the procedure at each period  $t$  to compute  $c_t$  as a function of the history  $h_{t-1}$ , and only then determine his action  $a_t$ . Thus, in particular, there cannot be deterministic calibrated procedures (because there are no leaky calibrated procedures); see Dawid (1985) and Oakes (1985).

In the case of smooth calibration, the procedure that we construct is deterministic, and thus leaky smoothly calibrated. However, there are also randomized leaky smoothly calibrated procedures. One example is the simple calibrated procedure of D. Foster (1999) in the one-dimensional case (where  $A = \{\text{“rain”}, \text{“no rain”}\}$  and  $C = [0, 1]$ ): the forecast there is “almost deterministic,” in the sense that the randomization is only between two very close forecasts (such as 29.99% and 30.01%), and so can be shown to be leaky smoothly calibrated. For another example, see footnote 22 in Section 4 below.

A particular instance of the leaky setup is one where the A-player uses a fixed reaction function  $g : C \rightarrow A$  that is a continuous mapping of forecasts to actions; thus,  $a_t = g(c_t)$  (independently of time  $t$  and history  $h_{t-1}$ ). In this case, leaky smooth calibration implies that most of the forecasts that are used must be approximate fixed points of  $g$ ; indeed, in every period in which the forecast is  $c$  the action is the same, namely,  $g(c)$ , and so the average of the actions in all the periods where the forecast is (close to)  $c$  is (close to)  $g(c)$  (use the continuity of  $g$  here); formally, see the arguments in part (iv) of the proof of Theorem 15 in Section 6). Thus, leaky procedures find (approximate) fixed points for arbitrary continuous functions  $g$ , and so must in general be more complex than the procedures that yield calibration (such as those obtained by Blackwell’s approachability); cf. the complexity class PPAD (Papadimitriou 1994) in the computer science literature (see also Hazan and Kakade 2012 for the connection to calibration).

## 2.4 Result

A strategy  $\sigma$  has *finite recall* and is *stationary* if there exists a finite integer  $R \geq 1$  and a function  $\tilde{\sigma} : (C \times A)^R \rightarrow C$  such that

$$\sigma(h_{T-1}) = \tilde{\sigma}(c_{T-R}, a_{T-R}, c_{T-R+1}, a_{T-R+1}, \dots, c_{T-1}, a_{T-1})$$

for every  $T > R$  and history  $h_{T-1} = (c_t, a_t)_{1 \leq t \leq T-1}$ . Thus, only the “window” consisting of the last  $R$  periods matters; the rest of the history, as well as the calendar time  $T$ , do not. Finally, a finite set  $D \subseteq C$  is a  $\delta$ -*grid* for  $C$  if for every  $c \in C$  there is  $d \equiv d(c) \in D$  such that  $\|d - c\| \leq \delta$ .

Our result is:

**Theorem 1** *For every  $\varepsilon > 0$  and  $L < \infty$  there is an  $(\varepsilon, L)$ -smoothly calibrated procedure. Moreover, the procedure may be taken to be:*

- *deterministic;*
- *leaky;*
- *with finite recall and stationary; and*
- *with all the forecasts lying on a finite grid.*<sup>7</sup>

The proof will proceed as follows. First, we construct deterministic finite-recall algorithms for the *online linear regression* problem (cf. Foster 1991, Azoury and Warmuth 2001); see Theorem 2 in Section 3. Next, we use these algorithms to get deterministic finite-recall *weakly calibrated* procedures (cf. Foster and Kakade 2004, 2006); see Theorem 10 in Section 4. Finally, we obtain smooth calibration from weak calibration; see Section 5.

## 3 Online Linear Regression

Classical linear regression tries to predict a variable  $y$  from a vector  $x$  of  $d$  variables (and so  $y \in \mathbb{R}$  and  $x \in \mathbb{R}^d$ ). There are observations  $(x_t, y_t)_t$ , and

---

<sup>7</sup>The sizes  $R$  of the recall and  $\delta$  of the grid depend on  $\varepsilon$ ,  $L$ , the dimension  $m$ , and the bound on the compact set  $C$ .

one typically assumes that<sup>8</sup>  $y_t = \theta'x_t + \epsilon_t$ , where  $\epsilon_t$  are (zero-mean normally distributed) error terms. The optimal estimator for  $\theta$  is then given by the least squares method; i.e.,  $\theta$  minimizes  $(1/T) \sum_{t=1}^T \psi_t(\theta)$  with

$$\psi_t(\theta) := (y_t - \theta'x_t)^2$$

for every  $t$ .

In the *online linear regression* problem (Foster 1991; see Section 1.1), the observations arrive sequentially, and at each time period  $t$  we want to determine  $\theta_t$  given the information at that time, namely,  $(x_1, y_1), \dots, (x_{t-1}, y_{t-1})$  and  $x_t$  only. The goal is to bound the difference between the mean square errors in the online case and the offline case (i.e., “in hindsight”); namely,

$$\frac{1}{T} \sum_{t=1}^T \psi_t(\theta_t) - \frac{1}{T} \sum_{t=1}^T \psi_t(\theta).$$

Thus, an online linear-regression *algorithm* takes as input a sequence  $(x_t, y_t)_{t \geq 1}$  in  $\mathbb{R}^d \times \mathbb{R}$  and gives as output a sequence  $(\theta_t)_{t \geq 1}$  in  $\mathbb{R}^d$ , such that  $\theta_t$  is a function only of  $x_1, y_1, \dots, x_{t-1}, y_{t-1}, x_t$ , for each  $t$ .

Our result is:

**Theorem 2** *Let  $X, Y$  be positive reals, and  $\varepsilon > 0$ . Then there exists a positive integer  $R_0 \equiv R_0(\varepsilon, X, Y, d)$  such that for every  $R > R_0$  there is an  $R$ -recall stationary deterministic algorithm that gives  $(\theta_t)_{t \geq 1}$ , such that*

$$\frac{1}{R} \sum_{t=T-R+1}^T [\psi_t(\theta_t) - \psi_t(\theta)] \leq \varepsilon(1 + \|\theta\|^2) \quad \text{and} \quad (4)$$

$$\frac{1}{T} \sum_{t=1}^T [\psi_t(\theta_t) - \psi_t(\theta)] \leq \varepsilon(1 + \|\theta\|^2) \quad (5)$$

*hold for every  $T \geq R$ , every  $\theta \in \mathbb{R}^d$ , and every sequence  $(x_t, y_t)_{t \geq 1}$  in  $\mathbb{R}^d \times \mathbb{R}$  with  $\|x_t\| \leq X$  and  $|y_t| \leq Y$  for all  $t$ .*

---

<sup>8</sup>Vectors are viewed as column vectors, and  $\theta'$  denotes the transpose of  $\theta$  (thus  $\theta'x$  is the scalar product  $\theta \cdot x$  of  $\theta$  and  $x$ ).

When in addition  $\theta$  is bounded,<sup>9</sup> say,  $\|\theta\| \leq M$ , the mean square error of our online algorithm is guaranteed not to exceed the optimal offline mean square error by more than  $\varepsilon(1 + M^2)$ .

The proof of Theorem 2 in the remainder of this section proceeds as follows.

- (i) *Forward algorithm* (Section 3.1). We start with the “forward algorithm” of Azoury and Warmuth (2001) and the resulting bound on the sum of regrets  $\psi_t(\theta_t) - \psi_t(\theta)$  (Theorem 3).
- (ii) *Discounted forward algorithm* (Section 3.2). We modify the procedure by introducing a  $\lambda$ -discount factor, which gives bounds on the discounted sum of regrets (Propositions 4 and 5).
- (iii) *Windowed discounted forward algorithm* (Section 3.3). We further modify the procedure by restricting the history to a window consisting only of the last  $R$  periods, which gives bounds on the sum of regrets over that window (Proposition 8).
- (iv) *Conclusion* (Section 3.3). One of the useful properties of discounting is that the weight of the initial segment from 1 up to  $T - R$  is small relative to the whole sum from 1 to  $T$ , and so dropping that initial segment has little effect on the procedure and the resulting estimates. We can thus choose an appropriate discount factor  $\lambda$  and a window size  $R$  yielding the desired bounds on the windowed sum of regrets, and thus also on the time average of the regrets (Proposition 9, which yields Theorem 2).

### 3.1 Forward Algorithm

The starting point is the following algorithm of Azoury and Warmuth (2001, Section 5.4). For each  $a > 0$ , the *a-forward algorithm* gives<sup>10</sup>  $\theta_t = Z_t^{-1}v_t$ ,

---

<sup>9</sup>For example,  $\theta$  lies in the unit simplex of  $\mathbb{R}^d$ .

<sup>10</sup> $Z_t^{-1}$  is the inverse of the  $d \times d$  matrix  $Z_t$  (which is invertible because  $a > 0$ ), and  $I$  denotes the identity matrix.

where

$$Z_t = aI + \sum_{q=1}^t x_q x_q' \quad \text{and} \quad v_t = \sum_{q=1}^{t-1} y_q x_q. \quad (6)$$

**Theorem 3 (Azoury and Warmuth 2001)** *For every  $a > 0$ , the  $a$ -forward algorithm yields*

$$\sum_{t=1}^T [\psi_t(\theta_t) - \psi_t(\theta)] \leq a \|\theta\|^2 + \sum_{t=1}^T y_t^2 \left( 1 - \frac{\det(Z_{t-1})}{\det(Z_t)} \right) \quad (7)$$

for every  $T \geq 1$ , every  $\theta \in \mathbb{R}^d$ , and every sequence  $(x_t, y_t)_{t \geq 1}$  in  $\mathbb{R}^d \times \mathbb{R}$ .

**Proof.** Theorem 5.6 and Lemma A.1 in Azoury and Warmuth (2001), where  $Z_t$  denotes their  $\eta_t^{-1}$  matrix; the second term in their formula (5.17) is non-negative since  $\eta_t$  is a positive definite matrix.<sup>11</sup>  $\square$

### 3.2 Discounted Forward Algorithm

Let  $a > 0$  and  $0 < \lambda < 1$ . The  $\lambda$ -discounted  $a$ -forward algorithm gives  $\theta_t = Z_t^{-1} v_t$ , where

$$Z_t = aI + \sum_{q=1}^t \lambda^{t-q} x_q x_q' \quad \text{and} \quad v_t = \sum_{q=1}^{t-1} \lambda^{t-q} y_q x_q. \quad (8)$$

**Proposition 4** *For every  $a > 0$  and  $0 < \lambda < 1$ , the  $\lambda$ -discounted  $a$ -forward algorithm yields*

$$\sum_{t=1}^T \lambda^{T-t} [\psi_t(\theta_t) - \psi_t(\theta)] \leq a \|\theta\|^2 + \sum_{t=1}^T \lambda^{T-t} y_t^2 \left( 1 - \lambda^d \frac{\det(Z_{t-1})}{\det(Z_t)} \right) \quad (9)$$

for every  $T \geq 1$ , every  $\theta \in \mathbb{R}^d$ , and every sequence  $(x_t, y_t)_{t \geq 1}$  in  $\mathbb{R}^d \times \mathbb{R}$ .

**Proof.** Let  $b := \sqrt{a(1-\lambda)}$ . From the sequence  $(x_t, y_t)_{t \geq 1}$  we construct a sequence  $(\tilde{x}_s, \tilde{y}_s)_{s \geq 1}$  in blocks as follows. For every  $t \geq 1$ , the  $t$ -th block  $B_t$  is

---

<sup>11</sup>Our statement is different from theirs because  $\psi_t$  equals twice  $L_t$ , and there is a misprinted sign in the first line of their formula (5.17).

of size  $d+1$  and consists of  $(\lambda^{-t/2}be^{(1)}, 0), \dots, (\lambda^{-t/2}be^{(d)}, 0), (\lambda^{-t/2}x_t, \lambda^{-t/2}y_t)$ , where  $e^{(i)}$  is the  $i$ -th unit vector in  $\mathbb{R}^d$ . The  $a$ -forward algorithm applied to  $(\tilde{x}_s, \tilde{y}_s)_{s \geq 1}$  yields the following.

For  $s = (d+1)t$ , i.e., at the end of the  $B_t$  block, we have<sup>12</sup>  $\sum_{s \in B_t} \tilde{x}_s \tilde{x}'_s = b^2 \lambda^{-t} \sum_{i=1}^d e^{(i)} (e^{(i)})' + \lambda^{-t} x_t x'_t = \lambda^{-t} (b^2 I + x_t x'_t)$ ; thus

$$\begin{aligned} \tilde{Z}_{(d+1)t} &= aI + \sum_{q=1}^t \sum_{s \in B_t} \tilde{x}_s \tilde{x}'_s = aI + \sum_{q=1}^t \lambda^{-q} (b^2 I + x_q x'_q) \\ &= \lambda^{-t} \left( aI + \sum_{q=1}^t \lambda^{t-q} x_q x'_q \right) = \lambda^{-t} Z_t \end{aligned}$$

(since  $\sum_{i=1}^d e^{(i)} (e^{(i)})' = I$  and  $b^2 = (1-\lambda)a$ ; recall (8)). Together with  $\tilde{v}_{(d+1)t} = \sum_{q=1}^t \sum_{s \in B_t} \tilde{y}_s \tilde{x}_s = \sum_{q=1}^t \lambda^{-q} y_q x_q = \lambda^{-t} v_t$  (only the first entry in each block has a nonzero  $\tilde{y}$ ), it follows that  $\tilde{\theta}_{(d+1)t}$  indeed equals  $\theta_t = Z_t^{-1} v_t$  as given by (8).

Next, for every  $t$  we have  $\sum_{s \in B_t} \tilde{\psi}_s(\tilde{\theta}_s) \geq \lambda^{-t} \psi_t(\theta_t)$  (all terms in the sum are nonnegative, and we drop all except the last one). Also, for every  $\theta \in \mathbb{R}^d$ ,

$$\sum_{s \in B_t} \tilde{\psi}_s(\theta) = \lambda^{-t} \left( b^2 \sum_{i=1}^d (\theta' e^{(i)})^2 + \psi_t(\theta) \right) = \lambda^{-t} (b^2 \|\theta\|^2 + \psi_t(\theta)).$$

Thus the left-hand side of (7) evaluated at the end of the  $T$ -th block  $B_T$  satisfies

$$\begin{aligned} LHS &\geq \sum_{t=1}^T \lambda^{-t} \psi_t(\theta_t) - a \|\theta\|^2 - b^2 \|\theta\|^2 \sum_{t=1}^T \lambda^{-t} - \sum_{t=1}^T \lambda^{-t} \psi_t(\theta) \\ &= \sum_{t=1}^T \lambda^{-t} [\psi_t(\theta_t) - \psi_t(\theta)] - \lambda^{-T} a \|\theta\|^2. \end{aligned}$$

On the right-hand side we get

$$RHS = \sum_{t=1}^T \sum_{s \in B_t} \tilde{y}_s^2 \left( 1 - \frac{\det(\tilde{Z}_{s-1})}{\det(\tilde{Z}_s)} \right) = \sum_{t=1}^T \lambda^{-t} y_t^2 \left( 1 - \frac{\det(\tilde{Z}_{(d+1)t-1})}{\det(\tilde{Z}_{(d+1)t})} \right)$$

---

<sup>12</sup>The notation  $\tilde{Z}_s, \tilde{\psi}_s, \dots$  pertains to the  $(\tilde{x}_s, \tilde{y}_s)_{s \geq 1}$  problem.

(again, only the last term in each block has nonzero  $\tilde{y}_s$ ). We have seen above that  $\tilde{Z}_{(d+1)t} = \lambda^{-t}Z_t$ ; thus  $\tilde{Z}_{(d+1)t-1} = \tilde{Z}_{(d+1)t} - \lambda^{-t}x_t x_t' = \lambda^{-t}(Z_t - x_t x_t') = \lambda^{-t+1}Z_{t-1} + (\lambda^{-t} - \lambda^{-t+1})aI$ . Therefore  $\det(\tilde{Z}_{(d+1)t-1}) \geq \det(\lambda^{-t+1}Z_{t-1})$  (indeed, if  $B$  is a positive definite matrix and  $\beta > 0$  then<sup>13</sup>  $\det(B + \beta I) > \det(B)$ ). Therefore we obtain

$$\begin{aligned} RHS &\leq \sum_{t=1}^T \lambda^{-t} y_t^2 \left( 1 - \frac{\det(\lambda^{-t+1}Z_{t-1})}{\det(\lambda^{-t}Z_t)} \right) \\ &= \sum_{t=1}^T \lambda^{-t} y_t^2 \left( 1 - \lambda^d \frac{\det(Z_{t-1})}{\det(Z_t)} \right) \end{aligned}$$

(the matrices  $Z_t$  are of size  $d \times d$ , and so  $\det(cZ_t) = c^d \det(Z_t)$ ). Recalling that  $LHS \leq RHS$  by (7) and multiplying by  $\lambda^T$  yields the result.  $\square$

**Remark.** From now on it will be convenient to assume that  $\|x_t\| \leq 1$  and  $|y_t| \leq 1$  (i.e.,  $X = Y = 1$ ); for general  $X$  and  $Y$ , multiply  $x_t, y_t, \theta_t, \psi_t, a$  by  $X, Y, Y/X, Y^2, X^2$ , respectively, in the appropriate formulas.

**Proposition 5** *For every  $a > 0$  and  $1/4 \leq \lambda < 1$  there exists a constant  $D_1 \equiv D_1(a, \lambda, d)$  such that the  $\lambda$ -discounted  $a$ -forward algorithm yields*

$$\sum_{t=1}^T \lambda^{T-t} [\psi_t(\theta_t) - \psi_t(\theta)] \leq a \|\theta\|^2 + D_1, \quad (10)$$

for every  $T \geq 1$ , every  $\theta \in \mathbb{R}^d$ , and every sequence  $(x_t, y_t)_{t \geq 1}$  in  $\mathbb{R}^d \times \mathbb{R}$  with  $\|x_t\| \leq 1$  and  $|y_t| \leq 1$  for all  $t$ .

**Proof.** Let  $K \geq 1$  be an integer such that  $1/4 \leq \lambda^K \leq 1/2$ . Given  $T \geq 1$ , let the integer  $m \geq 1$  satisfy  $(m-1)K < T \leq mK$ . Writing  $\zeta_t$  for  $\det(Z_t)$ ,

---

<sup>13</sup>Let  $\beta_1, \dots, \beta_d > 0$  be the eigenvalues of  $B$ ; then the eigenvalues of  $B + cI$  are  $\beta_1 + c, \dots, \beta_d + c$ , and so  $\det(B + cI) = \prod_i (\beta_i + c) > \prod_i \beta_i = \det(B)$ .

we have

$$\begin{aligned}
\sum_{t=1}^T \lambda^{T-t} \left(1 - \lambda^d \frac{\zeta_{t-1}}{\zeta_t}\right) &\leq \sum_{t=1}^{mK} \lambda^{T-t} \left(1 - \lambda^d \frac{\zeta_{t-1}}{\zeta_t}\right) \\
&= \lambda^T \sum_{j=1}^{m-1} \sum_{t=jK+1}^{(j+1)K} \lambda^{-t} \ln \left( \lambda^{-d} \frac{\zeta_t}{\zeta_{t-1}} \right) \\
&\leq \lambda^T \sum_{j=1}^{m-1} \lambda^{-(j+1)K} \sum_{t=jK+1}^{(j+1)K} \ln \left( \lambda^{-d} \frac{\zeta_t}{\zeta_{t-1}} \right) \\
&\leq \lambda^T \sum_{j=1}^{m-1} \lambda^{-(j+1)K} \ln \left( \lambda^{-dK} \frac{\zeta_{(j+1)K}}{\zeta_{jK}} \right) \quad (11)
\end{aligned}$$

(in the second line we have used  $1 - 1/u \leq \ln u$  for  $0 < u \leq 1$ , as in (4.21) in Azoury and Warmuth 2001; in the third line,  $\lambda^{-t} \leq \lambda^{-(j+1)K}$  since  $t \leq (j+1)K$  and  $\lambda < 1$ ).

Let  $B = (b_{ij})$  be a  $d \times d$  symmetric positive definite matrix with  $|b_{ij}| \leq \beta$  for all  $i, j$ , and let  $a > 0$ . Then  $a^d \leq \det(aI + B) \leq d!(a + \beta)^d$ . Indeed, the second inequality follows easily since the determinant is the sum of  $d!$  products of  $d$  elements each. For the first inequality, let  $\beta_1, \dots, \beta_d > 0$  be the eigenvalues of  $B$ ; then the eigenvalues of  $aI + B$  are  $a + \beta_1, \dots, a + \beta_d$ , and so  $\det(aI + B) = \prod_{i=1}^d (a + \beta_i) > a^d$ . Applying this to  $Z_t$  (using (8),  $|x_{t,i}x_{t,j}| \leq \|x_t\|^2 \leq 1$ , and  $\sum_{t=1}^T \lambda^{T-t} < 1/(1 - \lambda)$ ) yields

$$a^d \leq \zeta_t \equiv \det(Z_t) \leq d! \left( a + \frac{1}{1 - \lambda} \right)^d.$$

Therefore, since  $\lambda^{-K} \leq 4$ , we get

$$\lambda^{-dK} \frac{\zeta_{(j+1)K}}{\zeta_{jK}} \leq 4^d d! \left( 1 + \frac{1}{a(1 - \lambda)} \right)^d =: D,$$

and so (11) is

$$\begin{aligned} &\leq \lambda^T \sum_{j=1}^{m-1} (\lambda^{-K})^{j+1} \ln D \leq \lambda^T \frac{\lambda^{-K(m+1)} - \lambda^{-2K}}{\lambda^{-K} - 1} \ln D \\ &\leq \lambda^T \frac{\lambda^{-T} \lambda^{-K} - 0}{2 - 1} \ln D = 4 \ln D \end{aligned}$$

(since  $2 \leq \lambda^{-K} \leq 4$  and  $K(m+1) < T + K$ ). Substituting this in (9) and putting

$$D_1 := 4 \left( \ln d! + d \ln 4 + d \ln \left( 1 + \frac{1}{a(1-\lambda)} \right) \right) \quad (12)$$

completes the proof.  $\square$

### 3.3 Windowed Discounted Forward Algorithm

From now on it is convenient to put  $(x_t, y_t, \theta_t) = (0, 0, 0)$  for all  $t \leq 0$ .

Let  $a > 0$ ,  $0 < \lambda < 1$ , and integer  $R \geq 1$ . The  $R$ -windowed  $\lambda$ -discounted  $a$ -forward algorithm gives  $\theta_t = Z_t^{-1} v_t$ , where<sup>14</sup>

$$Z_t = aI + \sum_{q=t-R+1}^t \lambda^{t-q} x_q x_q' \quad \text{and} \quad v_t = \sum_{q=t-R+1}^{t-1} \lambda^{R-q} y_q x_q. \quad (13)$$

**Lemma 6** *For every  $a > 0$  and  $0 < \lambda < 1$  there exists a constant  $D_2 \equiv D_2(a, \lambda, d)$  such that if  $(\tilde{\theta}_t)_{t \geq 1}$  is given by the  $\lambda$ -discounted  $a$ -forward algorithm, and  $(\theta_t)_{t \geq 1}$  is given by the  $R$ -windowed  $\lambda$ -discounted  $a$ -forward algorithm for some integer  $R \geq 1$ , then*

$$\left| \psi_t(\tilde{\theta}_t) - \psi_t(\theta_t) \right| \leq D_2 \lambda^R \quad (14)$$

for every<sup>15</sup>  $t \geq 1$  and every sequence  $(x_t, y_t)_{t \geq 1}$  in  $\mathbb{R}^d \times \mathbb{R}$  with  $\|x_t\| \leq 1$  and  $|y_t| \leq 1$  for all  $t$ .

To prove this lemma we use the following basic result. The norm of a

<sup>14</sup>The sums below effectively start at  $\min\{t - R + 1, 1\}$  (because we put  $x_q = 0$  for  $q \leq 0$ ).

<sup>15</sup>For  $t \leq R$  we have  $\tilde{\theta}_t = \theta_t$  since they are given by the same formula.

matrix  $A$  is  $\|A\| := \max_{z \neq 0} \|Az\| / \|z\|$ .

**Lemma 7** For  $k = 1, 2$ , let  $c_k = A_k^{-1}b_k$ , where  $A_k$  is a  $d \times d$  symmetric matrix whose eigenvalues are all greater than or equal to some  $\alpha > 0$ , and  $\|b_k\| \leq M$ . Then  $\|c_k\| \leq M/\alpha$  and

$$\|c_1 - c_2\| \leq \frac{1}{\alpha} \|b_1 - b_2\| + \frac{M}{\alpha^2} \|A_1 - A_2\|.$$

**Proof.** First,  $\|c_k\| \leq \|A_k^{-1}\| \|b_k\| \leq (1/\alpha)M$  since  $\|A_k^{-1}\|$  equals the maximal eigenvalue of  $A_k^{-1}$ , which is the reciprocal of the minimal eigenvalue of  $A_k$ , and so  $\|A_k^{-1}\| \leq 1/\alpha$ .

Second, express  $c_1 - c_2$  as  $A_1^{-1}(b_1 - b_2) + A_1^{-1}(A_2 - A_1)A_2^{-1}b_2$ , to get

$$\|c_1 - c_2\| \leq \|A_1^{-1}\| \|b_1 - b_2\| + \|A_1^{-1}\| \|A_2 - A_1\| \|A_2^{-1}\| \|b_2\|$$

and the proof is complete.  $\square$

**Proof of Lemma 6.** For  $t \leq R$  we have  $\tilde{\theta}_t \equiv \theta_t$ , and so consider  $t > R$ . We have<sup>16</sup>  $\|\tilde{v}_t\|, \|v_t\| \leq \sum_{q=1}^{\infty} \lambda^q = 1/(1-\lambda)$ . The matrices  $\tilde{Z}_t$  and  $Z_t$  are the sum of  $aI$  and a positive-definite matrix, and so their eigenvalues are  $\geq a$ . Next,

$$\|\tilde{v}_t - v_t\| = \left\| \sum_{q=1}^{t-R} \lambda^{t-q} y_q x_q \right\| \leq \frac{\lambda^R}{1-\lambda};$$

similarly, for each element  $(\tilde{Z}_t - Z_t)_{ij}$  of  $\tilde{Z}_t - Z_t$  we have

$$\left| (\tilde{Z}_t - Z_t)_{ij} \right| = \left| \sum_{q=1}^{t-R} \lambda^{t-q} x_{q,i} x_{q,j} \right| \leq \frac{\lambda^R}{1-\lambda},$$

---

<sup>16</sup>Notation:  $\tilde{v}_t$  and  $\tilde{Z}_t$  pertain to the sequence  $\tilde{\theta}_t$  given by the  $\lambda$ -discounted  $a$ -forward algorithm, whereas  $v_t$  and  $Z_t$  pertain to the sequence  $\theta_t$  given by the  $R$ -windowed  $\lambda$ -discounted  $a$ -forward algorithm.

and so<sup>17</sup>  $\|\tilde{Z}_t - Z_t\| \leq d\lambda^R/(\lambda - 1)$ . Using Lemma 7 yields

$$\|\tilde{\theta}_t - \theta_t\| \leq \frac{1}{a} \frac{\lambda^R}{1 - \lambda} + \frac{1}{a^2} \frac{d\lambda^R}{1 - \lambda} = \frac{\lambda^R(a + d)}{(1 - \lambda)a^2}.$$

Hence

$$\begin{aligned} |\psi_t(\tilde{\theta}_t) - \psi_t(\theta_t)| &= |(y_t - \tilde{\theta}'_t x_t)^2 - (y_t - \theta'_t x_t)^2| \\ &= |(\tilde{\theta}'_t - \theta'_t)x_t \cdot (2y_t - (\tilde{\theta}'_t + \theta'_t)x_t)| \\ &\leq \|\tilde{\theta}_t - \theta_t\| \left(2 + \|\tilde{\theta}_t\| + \|\theta_t\|\right) \\ &\leq \frac{\lambda^R(a + d)}{(1 - \lambda)a^2} \left(2 + \frac{2}{(1 - \lambda)a}\right) = D_2\lambda^R, \end{aligned}$$

where

$$D_2 := \frac{2(a + d)(a(1 - \lambda) + 1)}{a^3(1 - \lambda)^2}; \quad (15)$$

this completes the proof.  $\square$

**Proposition 8** *For every  $a > 0$  and  $1/4 \leq \lambda < 1$  there exist constants  $D_1 \equiv D_1(a, \lambda, d)$  and  $D_2 \equiv D_2(a, \lambda, d)$  such that for every integer  $R \geq 1$  the  $R$ -windowed  $\lambda$ -discounted  $a$ -forward algorithm yields*

$$\begin{aligned} \frac{1}{R} \sum_{t=T-R+1}^T [\psi_t(\theta_t) - \psi_t(\theta)] &\leq (a\|\theta\|^2 + D_1) \left(1 - \lambda + \frac{\lambda}{R}\right) \\ &\quad + \frac{(\|\theta\| + 1)^2}{R(1 - \lambda)} + D_2\lambda^R \end{aligned} \quad (16)$$

for every  $T \geq 1$ , every  $\theta \in \mathbb{R}^d$ , and every sequence  $(x_t, y_t)_{t \geq 1}$  in  $\mathbb{R}^d \times \mathbb{R}$  with  $\|x_t\| \leq 1$  and  $|y_t| \leq 1$  for all  $t$ .

**Proof.** Let  $\tilde{\theta}_t$  be given by the  $\lambda$ -discounted  $a$ -forward algorithm. Put  $g_t := \psi_t(\theta_t) - \psi_t(\theta)$  (where  $\theta_t$  is given by the  $R$ -windowed  $\lambda$ -discounted  $a$ -forward algorithm) and  $\tilde{g}_t := \psi_t(\tilde{\theta}_t) - \psi_t(\theta)$ . Apply (10) at  $T$ , and also at each one

<sup>17</sup>Because  $\|A\| \leq d \max_{i,j} |a_{ij}|$  for any  $d \times d$  matrix  $A$ .

of  $T - R + 1, T - R + 2, \dots, T - 1$ ; multiply those by  $1 - \lambda$  and add them all up, to get

$$\begin{aligned} \sum_{t=1}^T \lambda^{T-t} \tilde{g}_t + (1 - \lambda) \sum_{r=1}^{R-1} \sum_{t=1}^{T-r} \lambda^{T-r-t} \tilde{g}_t &\leq (a \|\theta\|^2 + D_1)(1 + (R - 1)(1 - \lambda)) \\ &= (a \|\theta\|^2 + D_1)(R - R\lambda + \lambda). \end{aligned}$$

For  $t \leq T - R$ , the total coefficient of  $\tilde{g}_t$  on the left-hand side above is  $\lambda^{T-t} + (1 - \lambda) \sum_{r=1}^{R-1} \lambda^{T-r-t} = \lambda^{T-R+1-t}$ ; for  $T - R + 1 \leq t \leq T$ , it is  $\lambda^{T-t} + (1 - \lambda) \sum_{r=1}^{T-t} \lambda^{T-r-t} = 1$ . Therefore

$$\sum_{t=1}^{T-R} \lambda^{T-R+1-t} \tilde{g}_t + \sum_{t=T-R+1}^T \tilde{g}_t \leq (a \|\theta\|^2 + D_1)(R - R\lambda + \lambda).$$

Now  $\tilde{g}_t \geq -\psi_t(\theta) \geq -(\|\theta\| \|x_t\| + |y_t|)^2 \geq -(\|\theta\| + 1)^2$ , and so

$$\begin{aligned} \sum_{t=T-R+1}^T \tilde{g}_t &\leq (a \|\theta\|^2 + D_1)(R - R\lambda + \lambda) + (\|\theta\| + 1)^2 \sum_{t=1}^{T-R} \lambda^{T-R+1-t} \\ &\leq (a \|\theta\|^2 + D_1)(R - R\lambda + \lambda) + \frac{(\|\theta\| + 1)^2}{1 - \lambda}. \end{aligned}$$

Divide by  $R$  and use  $g_t \leq \tilde{g}_t + D_2 \lambda^R$  (by Proposition 6).  $\square$

Choosing appropriate  $\lambda$  and  $R$  allows us to bound the right-hand side of (16).

**Proposition 9** *For every  $\varepsilon > 0$  and  $a > 0$  there is  $\lambda_0 \equiv \lambda_0(\varepsilon, a, d) < 1$  such that for every  $\lambda_0 < \lambda < 1$  there is  $R_0 \equiv R_0(\varepsilon, a, d, \lambda) \geq 1$  such that for every  $R \geq R_0$  the  $R$ -windowed  $\lambda$ -discounted  $a$ -forward algorithm yields*

$$\frac{1}{R} \sum_{t=T-R+1}^T [\psi_t(\theta_t) - \psi_t(\theta)] \leq \varepsilon(1 + \|\theta\|^2), \quad (17)$$

$$\frac{1}{T} \sum_{t=1}^T [\psi_t(\theta_t) - \psi_t(\theta)] \leq \varepsilon(1 + \|\theta\|^2), \quad (18)$$

for every  $T \geq R$ , every  $\theta \in \mathbb{R}^d$ , and every sequence  $(x_t, y_t)_{t \geq 1}$  in  $\mathbb{R}^d \times \mathbb{R}$  with

$\|x_t\| \leq 1$  and  $|y_t| \leq 1$  for all  $t$ .

**Proof.** The right-hand side of (16) is

$$\leq \left( D_1(1 - \lambda) + \frac{D_1}{R} + \frac{2}{R(1 - \lambda)} + D_2\lambda^R \right) + \|\theta\|^2 \left( a(1 - \lambda) + \frac{a}{R} + \frac{2}{R(1 - \lambda)} \right)$$

(use  $\lambda/R \leq 1/R$  and  $(\|\theta\| + 1)^2 \leq 2\|\theta\|^2 + 2$ ). First, take  $1/4 \leq \lambda_0 < 1$  close enough to 1 so that  $a(1 - \lambda_0) \leq \varepsilon/4$  and  $D_1(a, \lambda_0, d) \cdot (1 - \lambda_0) \leq \varepsilon/4$  (recall formula (12) for  $D_1$  and use  $\lim_{x \rightarrow 0^+} x \ln x = 0$ ). Then, given  $\lambda \in [\lambda_0, 1)$ , take  $R_0 \geq 1$  large enough so that  $a/R_0 \leq \varepsilon/4$ ,  $D_1(a, \lambda, d)/R_0 \leq \varepsilon/4$ ,  $2/(R_0(1 - \lambda)) \leq \varepsilon/4$ , and  $D_2(a, \lambda, d)\lambda^{R_0} \leq \varepsilon/4$ . This shows (17) for every  $T \geq 1$ .

In particular, for  $T' < R$  we get  $(1/R) \sum_{t=1}^{T'} [\psi_t(\theta_t) - \psi_t(\theta)] \leq \varepsilon(1 + \|\theta\|^2)$  (because  $(x_t, y_t, \theta_t) = (0, 0, 0)$  for all  $t \leq 0$ ). For  $T \geq R$ , add up the inequalities (17) for the disjoint blocks of size  $R$  that end at  $t = T$ , together with the above inequality for the initial smaller block of size  $T' < R$  when  $T$  is not a multiple of  $R$ , to get  $(1/R) \sum_{t=1}^T [\psi_t(\theta_t) - \psi_t(\theta)] \leq \lceil T/R \rceil \varepsilon(1 + \|\theta\|^2) \leq 2(T/R)\varepsilon(1 + \|\theta\|^2)$ . Replacing  $\varepsilon$  with  $\varepsilon/2$  yields (18).  $\square$

**Remark.** Similar arguments show that, for  $\lambda_0 \leq \lambda < 1$ , the discounted average is also small:

$$\frac{1 - \lambda}{1 - \lambda^T} \sum_{t=1}^T \lambda^{T-t} [\psi_t(\theta_t) - \psi_t(\theta)] \leq \varepsilon(1 + \|\theta\|^2).$$

Proposition 9 yields the main result of this section, Theorem 2.

**Proof of Theorem 2.** Use Proposition 9 (with, say,  $a = 1$ ), and rescale everything by  $X$  and  $Y$  appropriately (see the Remark before Proposition 5).  $\square$

## 4 Weak Calibration

The notion of “weak calibration” was introduced by Kakade and Foster (2004) and Foster and Kakade (2006). The idea is as follows. Given a “test” function  $w : C \rightarrow \{0, 1\}$  that indicates which forecasts  $c$  to consider, let the corresponding score be<sup>18</sup>  $S_T^w := \left\| (1/T) \sum_{t=1}^T w(c_t)(a_t - c_t) \right\|$ . It can be shown that if  $S_T^w$  is small for every such  $w$ , then the calibration score  $K_T$  is also small.<sup>19</sup>

Now instead of the discontinuous indicator functions, *weak calibration* requires that  $S_T^w$  be small for Lipschitz continuous “weight” functions  $w : C \rightarrow [0, 1]$ . Specifically, let  $\varepsilon > 0$  and  $L < \infty$ . A procedure (i.e., a strategy of the C-player in the calibration game) is  $(\varepsilon, L)$ -*weakly calibrated* if there is  $T_0 \equiv T_0(\varepsilon, L)$  such that

$$S_T^w = \left\| \frac{1}{T} \sum_{t=1}^T w(c_t)(a_t - c_t) \right\| \leq \varepsilon \quad (19)$$

holds for every strategy of the A-player, every  $T > T_0$ , and every weight function  $w : C \rightarrow [0, 1]$  that is  $L$ -Lipschitz (i.e.,  $\mathcal{L}(w) \leq L$ ).

The importance of weak calibration is that, unlike regular calibration, it can be guaranteed by deterministic procedures (which are thus leaky): Kakade and Foster (2004) and Foster and Kakade (2006) have proven the existence of deterministic  $(\varepsilon, L)$ -weakly calibrated procedures. Moreover, as we will show in the next section, weak calibration is essentially equivalent to smooth calibration.

We now provide a deterministic  $(\varepsilon, L)$ -weakly calibrated procedure that in addition has finite recall and is stationary.

---

<sup>18</sup>The  $S_T$  scores are norms of averages, rather than averages of norms like the  $K_T$  scores. “Windowed” versions of the scores may also be considered (with the average taken over the last  $R$  periods only; cf. (4)).

<sup>19</sup>Specifically, if  $S_T^w \leq \varepsilon$  for all  $w : C \rightarrow \{0, 1\}$  then  $K_T \leq 2m\varepsilon$ . Indeed, for each coordinate  $i = 1, \dots, m$ , let  $C_+^i$  be the set of all  $c_t$  such that  $\bar{a}_{t,i} > c_{t,i}$ , and  $C_-^i$  the set of all  $c_t$  such that  $\bar{a}_{t,i} < c_{t,i}$ . Taking  $w$  to be the indicator of  $C_+^i$  yields  $S_T^w = (1/T) \sum_t [\bar{a}_{t,i} - c_{t,i}]_+ \leq \varepsilon$  (where  $[z]_+ := \max\{z, 0\}$ ); similarly, the indicator of  $C_-^i$  yields  $(1/T) \sum_t [\bar{a}_{t,i} - c_{t,i}]_- \leq \varepsilon$ . Adding the two inequalities gives  $(1/T) \sum_t |\bar{a}_{t,i} - c_{t,i}| \leq 2\varepsilon$ . Since this holds for each one of the  $m$  coordinates, it follows that  $K_T \leq 2m\varepsilon$ .

**Theorem 10** *For every  $\varepsilon > 0$  and  $L < \infty$  there exists an  $(\varepsilon, L)$ -weakly calibrated deterministic procedure that has finite recall and is stationary; moreover, all its forecasts may be taken to lie on a finite grid.*

The proof uses the result of Theorem 2. The basic idea is to use the forecast itself as part of the input to the forecast—which adds a fixed-point construct to the regression. Assume for starters that  $a$  and  $c$  are one-dimensional, and also that we have only a single weight function  $w$ . Consider the online linear regression problem with  $x_t = (c_t, w(c_t))$  and  $y_t = a_t$ . Given the history  $h_{t-1}$  up to and including time  $t - 1$ , if we knew the value of  $c_t$  then we would get a forecast  $\hat{a}_t := \theta'_t x_t$  for which, by equation (5), the regret is small. But we do not know  $c_t$ , as it is going in fact to be our forecast: that is, we want to choose  $c_t$  so that the resulting  $\hat{a}_t$  satisfies  $\hat{a}_t = c_t$ . This requires solving a fixed-point problem (which is possible since the mapping  $H$  from  $c_t$  to  $\hat{a}_t$  is continuous), and indeed yields the desired  $c_t$ . Now equation (5) yields, for an appropriate  $\varepsilon > 0$ ,

$$\frac{1}{T} \sum_{t=1}^T (a_t - \hat{a}_t)^2 \leq \frac{1}{T} \sum_{t=1}^T (a_t - \theta'_t x_t)^2 + \varepsilon$$

for all  $\theta$ . But  $\hat{a}_t = c_t$ , and so taking  $\theta = (1, \sqrt{\varepsilon})$  gives

$$\begin{aligned} \frac{1}{T} \sum_{t=1}^T (a_t - c_t)^2 &\leq \frac{1}{T} \sum_{t=1}^T (a_t - c_t - \sqrt{\varepsilon} w(c_t))^2 + \varepsilon \\ &\leq \frac{1}{T} \sum_{t=1}^T (a_t - c_t)^2 - 2\sqrt{\varepsilon} \frac{1}{T} \sum_{t=1}^T w(c_t)(a_t - c_t) + \varepsilon + \varepsilon \end{aligned}$$

(in the second line we have used  $w(c_t) \in [0, 1]$ ). Therefore

$$\frac{1}{T} \sum_{t=1}^T w(c_t)(a_t - c_t) \leq \sqrt{\varepsilon};$$

together with the similar computation for  $\theta = (1, -\sqrt{\varepsilon})$  we get

$$S_T^w = \left| \frac{1}{T} \sum_{t=1}^T w(c_t)(a_t - c_t) \right| \leq \sqrt{\varepsilon},$$

as desired. To deal with  $m$ -dimensional  $a$  and  $c$  we use  $m$  separate online regressions, one for each coordinate; to deal with all the  $L$ -Lipschitz weight functions  $w$ , we take an appropriate finite grid.

**Proof.** (i) *Preliminaries.* Without loss of generality assume that  $A \subseteq C \subseteq [0, 1]^m$  (one can always translate the sets  $A$  and  $C$ —which does not affect (19)—and rescale them—which just rescales the Lipschitz constant); assume also that  $L \geq 1$  (as  $L$  increases there are more Lipschitz functions) and  $\varepsilon \leq 1$ .

For every  $b \in \mathbb{R}^m$  let  $\gamma(b) := \arg \min_{c \in C} \|c - b\|$  be the closest point to  $b$  in  $C$  (it is well defined and unique since  $C$  is a convex compact set); then

$$\|c - b\| \geq \|c - \gamma(b)\| \tag{20}$$

for every  $c \in C$  (because

$$\|c - b\|^2 = \|c - \gamma(b)\|^2 + \|b - \gamma(b)\|^2 - 2(b - \gamma(b)) \cdot (c - \gamma(b))$$

and the third term is  $\leq 0$ ); moreover,  $\gamma(b) = b$  when  $b \in C$ .

Let  $\varepsilon_1 := \varepsilon/(2\sqrt{m})$ . Denote by  $W_L$  the set of weight functions  $w : C \rightarrow [0, 1]$  with  $\mathcal{L}(w) \leq L$ . By Lemma 18 in the Appendix, there exist  $d$  functions  $f_1, \dots, f_d$  in  $W_L$  such that for every  $w \in W_L$  there is a vector  $\varpi \equiv \varpi_w \in [0, 1]^d$  with<sup>20</sup>

$$\max_{c \in C} \left| w(c) - \sum_{i=1}^d \varpi_i f_i(c) \right| \leq \varepsilon_1. \tag{21}$$

Denote  $F(c) := (f_1(c), \dots, f_d(c)) \in [0, 1]^d$ ; thus  $\|F(c)\| \leq \sqrt{d}$ . Without loss of generality we assume that the set  $\{f_1, \dots, f_d\}$  includes the “ $j$ -th coordinate function,” which maps each  $c \in C$  to its  $j$ -th coordinate  $c_j$ ; say,  $f_j(c) = c_j$

---

<sup>20</sup>Since  $W_L$  is compact in the sup norm, there are  $f_1, \dots, f_d \in W_L$  such that for every  $w \in W_L$  there is  $1 \leq i \leq d$  with  $\max_{c \in C} |w(c) - f_i(c)| \leq \varepsilon_1$ . Lemma 18 improves on this, in getting a much smaller  $d$  by using linear combinations with bounded coefficients.

for  $j = 1, \dots, m$  (thus  $d > m$ ; in fact  $d$  is much larger than  $m$ ).

Let  $\varepsilon_2 := \varepsilon / (m + m(1 + d)^2 + d^2)$  (where  $d$  is given by Lemma 18 in the Appendix, and depends on  $\varepsilon, m$ , and  $L$ ) and  $\varepsilon_3 := (\varepsilon_2)^2$ .

(ii) *The function  $H$ .* Let  $\lambda$  and  $R$  be given by Theorem 2 and Proposition 9 for  $a = 1$ ,  $X = \sqrt{d}$ ,  $Y = 1$ , and  $\varepsilon = \varepsilon_3$ . For each  $j = 1, \dots, m$  consider the sequence  $(x_t, y_t^{(j)})_{t \geq 1} = (F(c_t), a_{t,j})_{t \geq 1}$  in  $\mathbb{R}^d \times \mathbb{R}$ , where  $a_t \in A$  is determined by the A-player, and  $c_t \in C$  is constructed inductively as follows.

Let the history be  $h_{t-1} = (c_1, a_1, \dots, c_{t-1}, a_{t-1})$ . For each  $c \in \mathbb{R}^m$ , let<sup>21</sup>

$$\begin{aligned} Z_t(c) &= I + \sum_{q=1}^{R-1} \lambda^{R-q} x_q x_q' + F(c)F(c)' \in \mathbb{R}^{d \times d}, \\ v_t^{(j)} &= \sum_{q=1}^{R-1} \lambda^{R-q} a_{q,j} x_q \in \mathbb{R}^d, \\ H_{t,j}(c) &= \left( Z_t(c)^{-1} v_t^{(j)} \right)' F(c) \in \mathbb{R}, \\ H_t(c) &= (H_{t,1}(c), \dots, H_{t,m}(c)) \in \mathbb{R}^m \end{aligned}$$

(where  $x_q = F(c_q)$  for  $q < t$ ). Finally, we extend the function  $H_t$  to all of  $\mathbb{R}^m$  by putting  $H_t(b) := H_t(\gamma(b))$  for every  $b \in \mathbb{R}^m$ ; i.e., we project  $b$  to its closest point  $\gamma(b)$  in  $C$ , and then apply  $H_t$  to it.

(iii) *Fixed point of  $H$ .* For every  $c \in C$  we have  $\|v_t^{(j)}\| \leq \sqrt{d}\lambda/(1 - \lambda)$  (since  $|a_{q,j}| \leq 1$  and  $\|x_q\| = \|F(c_q)\| \leq \sqrt{d}$ ), and so  $\|Z_t(c)^{-1} v_t^{(j)}\| \leq \sqrt{d}\lambda/(1 - \lambda)$  by Lemma 7 ( $Z_t(c)$  is positive definite and its eigenvalues are  $\geq 1$ ), which finally implies that  $|H_{t,j}(c)| \leq \sqrt{d}\lambda/(1 - \lambda) \cdot \sqrt{d} = d\lambda/(1 - \lambda) =: K$ . Therefore the restriction of  $H_t$  to the compact and convex set  $[-K, K]^m$ , which is clearly a continuous function (since, again,  $Z_t(c)$  is positive definite and its eigenvalues are  $\geq 1$ ), has a fixed point (by Brouwer's fixed-point theorem), which we denote  $b_t$  (any fixed point will do);<sup>22</sup> put  $c_t := \gamma(b_t) \in C$ .

<sup>21</sup>A subscript  $j$  stands for the  $j$ -th coordinate (e.g.,  $a_{t,j}$  is the  $j$ -th coordinate of  $a_t$ ), whereas a superscript  $j$  refers to the  $j$ -th procedure (e.g.,  $v_t^{(j)}$ ).

<sup>22</sup>There may be more than one fixed point here, in which case we may choose the fixed point at random, and obtain a *randomized* procedure that satisfies everything the deterministic procedure does. Using it yields in Theorem 1 a randomized procedure that is *leaky* smoothly calibrated (cf. Section 2.3).

Thus

$$c_t = \gamma(b_t) \quad \text{and} \quad b_t = H_t(b_t) = H_t(c_t).$$

Define  $x_t := F(\gamma(b_t)) = F(c_t)$  and  $\theta_t^{(j)} := Z_t(c_t)^{-1}v_t^{(j)} \in \mathbb{R}^d$ . Then  $Z_t(c_t) = I + \sum_{q=1}^R \lambda^{R-q} x_q x_q'$ , and thus it corresponds to the  $R$ -windowed  $\lambda$ -discounted 1-forward algorithm (see (13)). Therefore, for every  $j = 1, \dots, m$  and every  $\theta^{(j)} \in \mathbb{R}^d$  we have by (5)

$$\frac{1}{T} \sum_{t=1}^T \left[ \psi_t^{(j)}(\theta_t^{(j)}) - \psi_t^{(j)}(\theta^{(j)}) \right] \leq \varepsilon_3 \left( 1 + \left\| \theta^{(j)} \right\|^2 \right) \quad (22)$$

for all  $T \geq T_0 \equiv R$ , where  $\psi_t^{(j)}(\theta) = (a_{t,j} - \theta' x_t)^2$ , and thus  $\psi_t^{(j)}(\theta_t^{(j)}) = (a_{t,j} - b_{t,j})^2$  (recall that  $b_{t,j} = H_{t,j}(b_t) = \left( \theta_t^{(j)} \right)' F(\gamma(b_t)) = \left( \theta_t^{(j)} \right)' x_t$ ). Summing over  $j$  yields

$$\frac{1}{T} \sum_{t=1}^T \sum_{j=1}^m \left[ \psi_t^{(j)}(\theta_t^{(j)}) - \psi_t^{(j)}(\theta^{(j)}) \right] \leq \varepsilon_3 \left( m + \sum_{j=1}^m \left\| \theta^{(j)} \right\|^2 \right).$$

Now  $\sum_{j=1}^m \psi_t^{(j)}(\theta_t^{(j)}) = \sum_{j=1}^m (a_{t,j} - b_{t,j})^2 = \|a_t - b_t\|^2 \geq \|a_t - \gamma(b_t)\|^2 = \|a_t - c_t\|^2 = \sum_{j=1}^m (a_{t,j} - c_{t,j})^2$  (by the definition of  $\gamma(b_t)$  and (20), since  $a_t \in A \subseteq C$ ), and therefore

$$\frac{1}{T} \sum_{t=1}^T \sum_{j=1}^m \left[ (a_{t,j} - c_{t,j})^2 - \psi_t^{(j)}(\theta^{(j)}) \right] \leq \varepsilon_3 \left( m + \sum_{j=1}^m \left\| \theta^{(j)} \right\|^2 \right). \quad (23)$$

(iv) *Estimating  $S_T^w$ .* Given a weight function  $w \in W_L$ , let the vector  $\varpi \equiv \varpi_w \in [0, 1]^d$  satisfy (21), i.e.,  $|w(c) - \varpi' F(c)| \leq \varepsilon_1$  for all  $c \in C$ . Take  $u = (u_j)_{j=1, \dots, m} \in \mathbb{R}^m$  with  $\|u\| = 1$ . For every  $j = 1, \dots, m$ , take  $\theta^{(j)} = e^{(j)} + \varepsilon_2 u_j \varpi \in \mathbb{R}^d$ , where  $e^{(j)} \in \mathbb{R}^d$  is the  $j$ -th unit vector; thus  $\left\| \theta^{(j)} \right\| \leq 1 + \varepsilon_2 d \leq 1 + d$  (since  $\varepsilon_2 \leq \varepsilon \leq 1$ ). We have

$$(\theta^{(j)})' x_t = (\theta^{(j)})' F(c_t) = c_{t,j} + \varepsilon_2 (\varpi' F(c_t)) u_j$$

(since  $f_j(c) = c_j$  for  $j \leq m$ ), and hence

$$\begin{aligned} (a_{t,j} - c_{t,j})^2 - \psi_t^{(j)}(\theta^{(j)}) &= (a_{t,j} - c_{t,j})^2 - (a_{t,j} - c_{t,j} - \varepsilon_2 (\varpi' F(c_t)) u_j)^2 \\ &= 2\varepsilon_2 (\varpi' F(c_t)) u_j (a_{t,j} - c_{t,j}) - (\varepsilon_2 \varpi' F(c_t))^2 u_j^2. \end{aligned}$$

Summing over  $j = 1, \dots, m$  yields

$$\begin{aligned} \sum_{j=1}^m \left[ (a_{t,j} - c_{t,j})^2 - \psi_t^{(j)}(\theta^{(j)}) \right] &= 2\varepsilon_2 (\varpi' F(c_t)) u'(a_t - c_t) - (\varepsilon_2 \varpi' F(c_t))^2 \|u\|^2 \\ &\geq 2\varepsilon_2 (\varpi' F(c_t)) u'(a_t - c_t) - (\varepsilon_2)^2 d^2 \\ &\geq 2\varepsilon_2 w(c_t) u'(a_t - c_t) - \varepsilon_1 \cdot 2\varepsilon_2 \|u\| \|a_t - c_t\| - (\varepsilon_2)^2 d^2 \\ &\geq 2\varepsilon_2 w(c_t) u'(a_t - c_t) - 2\varepsilon_1 \varepsilon_2 \sqrt{m} - (\varepsilon_2)^2 d^2 \end{aligned}$$

(since:  $\|u\| = 1$ ,  $|\varpi' F(c)| \leq d$  [the coordinates of  $\varpi$  are between  $-1$  and  $1$  and those of  $F(c)$  between  $0$  and  $1$ ],  $\|a_t - c_t\| \leq \sqrt{m}$  (since  $a_t, c_t \in [0, 1]^m$ , and recall (21)).

Together with (23) we get (recall that  $\varepsilon_3 = (\varepsilon_2)^2$  and  $\varepsilon_1 = \varepsilon/(2\sqrt{m})$ ):

$$2\varepsilon_2 \cdot \frac{1}{T} \sum_{t=1}^T w(c_t) u'(a_t - c_t) \leq (\varepsilon_2)^2 (m + m(1 + d)^2 + d^2) + \varepsilon \varepsilon_2;$$

hence, dividing by  $2\varepsilon_2$  and recalling that  $\varepsilon_2 = \varepsilon/(m + m(1 + d)^2 + d^2)$ :

$$u \cdot \frac{1}{T} \sum_{t=1}^T w(c_t) (a_t - c_t) \leq \frac{\varepsilon}{2} + \frac{\varepsilon}{2} = \varepsilon.$$

Since  $u \in \mathbb{R}^m$  with  $\|u\| = 1$  was arbitrary, the proof of (19) is complete.

(v) *Grid.* For the “moreover” statement, let  $\varepsilon_4 := \varepsilon/(L\sqrt{m} + 1)$ , and take  $D \subseteq C$  to be a finite  $\varepsilon_4$ -grid in  $C$ ; i.e., for every  $c \in C$  there is  $d(c) \in D$  with  $\|d(c) - c\| \leq \varepsilon_4$ . Replace the forecast  $c_T$  obtained above with  $\tilde{c}_T := d(c_T)$ ; then, for every  $a_T \in A$ , we have

$$\|w(c_T)(a_T - c_T) - w(\tilde{c}_T)(a_T - \tilde{c}_T)\| \leq L\varepsilon_4 \sqrt{m} + \varepsilon_4 = \varepsilon.$$

Therefore the score  $S_T^w$  changes by at most  $\varepsilon$ , and so it is at most  $2\varepsilon$ .  $\square$

## 5 Smooth Calibration

In this section we show (Propositions 13 and 14) that weak calibration and smooth calibration are essentially equivalent (albeit with different constants  $\varepsilon, L$ ). The existence of weakly calibrated procedures (Theorem 10, proved in the previous section) then implies the existence of smoothly calibrated procedures, which proves Theorem 1.

We first show how to go from weak to smooth calibration. When comparing the two scores, we see that the smooth calibration score uses weighted averages rather than sums:  $\sum_{s=1}^T \Lambda(c_s, c_t)(a_s - c_s)$  is divided by  $\sum_{s=1}^T \Lambda(c_s, c_t)$ . The following useful lemma shows how to bound the latter using the former.

**Lemma 11** *There exists a constant  $0 < \gamma \equiv \gamma_C < \infty$  that depends only on the dimension  $m$  and the diameter  $\alpha$  of  $C$  such that for any  $L$ -smoothing weight function  $\Lambda$ , any  $c_1, \dots, c_T \in C$ , and any<sup>23</sup>  $b_1, \dots, b_T \in C - C$ , putting*

$$B_t := \sum_{s=1}^T \Lambda(c_s, c_t) b_s \quad \text{and} \quad W_t := \sum_{s=1}^T \Lambda(c_s, c_t)$$

for all  $1 \leq t \leq T$ , we have

$$\frac{1}{T} \sum_{t=1}^T \frac{\|B_t\|}{W_t} \leq \gamma L^{m/2} \left( \frac{1}{T} \max_{1 \leq t \leq T} \|B_t\| \right)^{1/2}.$$

**Proof.** Let  $D \subset \mathbb{R}^m$  be an  $m$ -dimensional cube with sides of length  $\alpha$  that contains  $C$  (such a cube exists because the diameter of  $C$  is  $\alpha$ ). Let  $D_1, \dots, D_M$  be a partition of  $D$  into disjoint cubes with sides of length  $1/(2L\sqrt{m})$ ; the diameter of each such cube is thus  $1/(2L)$ , and the number of cubes is  $M = \lceil 2\alpha L\sqrt{m} \rceil^m$ .

---

<sup>23</sup>The set  $C - C$  consists of all  $b = b' - b''$  with  $b', b'' \in C$ .

Put  $\kappa := (1/T) \max_{1 \leq t \leq T} \|B_t\|$  and  $\bar{b}_t \equiv \bar{b}_t^\Lambda := B_t/W_t$ . When  $W_t$  is large, the inequality  $\|B_t\|/W_t \leq \kappa T/W_t$  provides a good bound; we will show that, for a large proportion of indices  $t$ , this is indeed the case.

Given  $\eta > 0$  (which will be specified later), call a cube  $D_j$  *good* if it contains at least  $\eta T$  elements of the sequence  $c_1, \dots, c_T$  (i.e.,  $|\{t \leq T : c_t \in D_j\}| \geq \eta T$ ), and *bad* otherwise; call an index  $t \leq T$  *good* if  $c_t$  belongs to some good cube  $D_i$ , and *bad* otherwise.

If  $c_t$  and  $c_s$  belong to the same cube  $D_j$  then  $\|c_s - c_t\| \leq \text{diam}(D_j) = 1/(2L)$ , and so  $1 - \Lambda(c_s, c_t) = \Lambda(c_t, c_t) - \Lambda(c_s, c_t) \leq L\|c_s - c_t\| \leq 1/2$ , which implies that  $\Lambda(c_s, c_t) \geq 1/2$ . Therefore for every good  $t$  we have  $W_t = \sum_{s \leq T} \Lambda(c_s, c_t) \geq (1/2)\eta T$ , and thus  $\|\bar{b}_t\| \leq 2T\kappa/(\eta T) = 2\kappa/\eta$ , which then gives

$$\frac{1}{T} \sum_{\text{good } t \leq T} \|\bar{b}_t\| \leq \frac{2\kappa}{\eta}.$$

The number of bad  $t$  is less than  $M \cdot \eta T$  (because each bad cube contains less than  $\eta T$  elements of  $c_1, \dots, c_T$ , and there are  $M$  cubes). For every  $s$  we have  $\|b_s\| \leq \text{diam}(C) = \alpha$ , and so  $\bar{b}_t$ , which is a weighted average of the  $b_s$ , satisfies  $\|\bar{b}_t\| \leq \alpha$  as well. Thus

$$\frac{1}{T} \sum_{\text{bad } t \leq T} \|\bar{b}_t\| \leq \frac{1}{T} M \cdot \eta T \cdot \alpha = \alpha \eta M.$$

Adding the last two displayed inequalities and choosing  $\eta = \sqrt{(2\kappa)/(\alpha M)}$  yields

$$\frac{1}{T} \sum_{t \leq T} \|\bar{b}_t\| \leq 2\sqrt{2\kappa\alpha M};$$

recalling that  $M = \lceil 2\alpha L\sqrt{m} \rceil^m$  gives the result, with  $\gamma$  essentially equal to  $2^{(m+3)/2} \alpha^{m/2} m^{m/4}$ .  $\square$

An immediate consequence is that smooth calibration is a weaker requirement than calibration.

**Corollary 12** *Calibration implies smooth calibration.*

**Proof.** We will show that<sup>24</sup>  $K_T^\Lambda = O(\sqrt{K_T})$  for each fixed  $L$  (where  $K_T^\Lambda$  is the  $\Lambda$ -smoothly calibrated score for any  $L$ -Lipschitz  $\Lambda$ , and  $K_T$  is the regular calibration score). Indeed, for every  $c_t \in C$  we have (use  $0 \leq \Lambda(\cdot, \cdot) \leq 1$ )

$$\begin{aligned} \frac{1}{T} \left\| \sum_{s=1}^T \Lambda(c_s, c_t)(a_s - c_s) \right\| &\leq \frac{1}{T} \sum_{s=1}^T \Lambda(c_s, c_t) \|a_s - c_s\| \\ &\leq \frac{1}{T} \sum_{s=1}^T \|a_s - c_s\| = K_T. \end{aligned}$$

Now apply Lemma 11 with  $b_s = a_s - c_s$  for all  $s$ .  $\square$

Returning to our proof, we can finally obtain smoothly calibrated procedures with the desired properties, proving Theorem 1.

**Proposition 13** *An  $(\varepsilon, L)$ -weakly calibrated procedure is  $(\varepsilon', L)$ -smoothly calibrated for  $\varepsilon' = \gamma L^{m/2} \varepsilon^{1/2}$  (with the constant  $\gamma \equiv \gamma_C$  given by Lemma 11).*

**Proof.** For any  $L$ -Lipschitz smoothing function  $\Lambda$ , Lemma 11 with  $b_s = a_s - c_s$  for all  $s$  yields

$$K_T^\Lambda \leq \gamma L^{m/2} \left( \sup_{w \in W_L} S_T^w \right)^{1/2}$$

(because  $\Lambda(\cdot, c_t) \in W_L$  for all  $c_t$ ; recall the definition (19) of  $S_T^w$  in Section 4). Therefore  $\sup_{w \in W_L} S_T^w \leq \varepsilon$  implies  $K_T^\Lambda \leq \gamma L^{m/2} \varepsilon^{1/2} = \varepsilon'$   $\square$

**Proof of Theorem 1.** Apply Theorem 10 and Proposition 13, and recall (Section 2.3) that for deterministic procedures leaks do not matter.  $\square$

As an aside, we now show how to go from smooth to weak calibration.

**Proposition 14** *An  $(\varepsilon, L)$ -smoothly calibrated procedure is  $(\varepsilon', L')$ -weakly calibrated, where<sup>25</sup>  $\varepsilon' = \Omega(\varepsilon^{1/2} L^{m/2})$  and  $L' = O(\varepsilon^{1/2} L^{(m+2)/2})$ .*

<sup>24</sup>The notations  $f(x) = O(g(x))$ ,  $f(x) = \Omega(g(x))$ , and  $f(x) = \Theta(g(x))$  mean, as usual, that there are constants  $C < \infty$  and  $c > 0$  such that for all  $x$  we have, respectively,  $f(x) \leq Cg(x)$ ,  $f(x) \geq cg(x)$ , and  $cg(x) \leq f(x) \leq Cg(x)$ . In our case  $x$  stands for  $(\varepsilon, L)$ ; the dimension  $m$  is assumed fixed.

<sup>25</sup>We have not tried to optimize the estimates for  $\varepsilon'$  and  $L'$ .

**Proof.** Let  $(D_j)_{j=1,\dots,M}$  be a partition of  $[0, 1]^m \supseteq C$  into disjoint cubes with sides  $\delta := 1/(L\sqrt{m})$ ; the diameter of each cube is thus  $\delta\sqrt{m} = 1/L$ , and the number of cubes is  $M = \delta^{-m} = L^m m^{m/2}$ . Let  $\varepsilon_1 := \sqrt{\varepsilon L^m}$  and  $\varepsilon_2 := \varepsilon_1 m^{m/2}/M = \sqrt{\varepsilon/L^m}$ . Take  $L' := L\varepsilon_1/2 = \sqrt{\varepsilon L^{m+2}}/2$  and  $\varepsilon' = \varepsilon_1(1 + \sqrt{m} + \sqrt{m^{m+1}}) = \sqrt{\varepsilon L^m}(1 + \sqrt{m} + \sqrt{m^{m+1}})$ .

Fix  $a_t, c_t \in C \subseteq [0, 1]^m$  for  $t = 1, \dots, T$ , and a weight function  $w$  in  $W_{L'}$ . Assume that  $K_T^\Lambda \leq \varepsilon$  holds for every smoothing function  $\Lambda$  that is  $L$ -Lipschitz in the first coordinate; we will show that  $S_T^w \leq \varepsilon'$  (where  $K_T^\Lambda$  and  $S_T^w$  are given by (2) and (19), respectively).

Let  $V \subseteq \{1, \dots, T\}$  be the set of indices  $t$  such that the cube  $D_j$  that contains  $c_t$  includes at least a fraction  $\varepsilon_2$  of  $c_1, \dots, c_T$ , i.e.,  $|\{s \leq T : c_s \in D_j\}| \geq \varepsilon_2 T$ . Then

$$T - |V| = |\{t \leq T : t \notin V\}| < M \cdot \varepsilon_2 T = \varepsilon_2 M T, \quad (24)$$

because there are at most  $M$  cubes containing less than  $\varepsilon_2 T$  points each.

We distinguish two cases.

*Case 1:*  $\max_{t \in V} w(c_t) < \varepsilon_1$ . Since  $\|a_t - c_t\| \leq \sqrt{m}$ , we have  $\|\sum_{t \in V} w(c_t)(a_t - c_t)\| \leq |V| \cdot \varepsilon_1 \cdot \sqrt{m} \leq \varepsilon_1 \sqrt{m} T$  (use  $|V| \leq T$ ), and  $\|\sum_{t \notin V} w(c_t)(a_t - c_t)\| \leq (T - |V|) \cdot 1 \cdot \sqrt{m} = \varepsilon_2 M \sqrt{m} T$  (use (24)). Adding and dividing by  $T$  yields

$$S_T^w \leq (\varepsilon_1 + \varepsilon_2 M) \sqrt{m} = \varepsilon_1 \sqrt{m} (1 + m^{m/2}) < K \varepsilon_1 = \varepsilon'.$$

*Case 2:*  $\max_{t \in V} w(c_t) \geq \varepsilon_1$ . Let  $s \in V$  be such that  $w(c_s) = \max_{t \in V} w(c_t) \geq \varepsilon_1$ , and let  $R \subseteq V$  be the set of indices  $r$  such that  $c_r$  lies in the same cube  $D_j$  as  $c_s$ .

For each  $r$  in  $R$ , proceed as follows. First, we have  $|w(c_s) - w(c_r)| \leq L' \|c_s - c_r\| \leq L' \cdot \delta \sqrt{m} = L\varepsilon_1/2 \cdot (1/L) = \varepsilon_1/2$ , and so

$$w(c_r) \geq w(c_s) - \frac{\varepsilon_1}{2} \geq \varepsilon_1 - \frac{\varepsilon_1}{2} = \frac{\varepsilon_1}{2}. \quad (25)$$

Next, put  $w^r(c) := \min\{w(c), w(c_r)\}$  and  $\Lambda(c, c_r) := w^r(c)/w(c_r)$  for  $r \in R$  (and, for  $t \notin R$ , put, say,  $\Lambda(c, c_t) = 1$  for all  $c$ ); then  $\mathcal{L}(\Lambda(\cdot, c_r)) \leq$

$\mathcal{L}(w)/w(c_r) \leq L' / (\varepsilon_1/2) = L$ , and so, by our assumption

$$\frac{1}{T} \sum_{r \in R} \|\bar{a}_r^\Lambda - c_r^\Lambda\| \leq \frac{1}{T} \sum_{t \leq T} \|\bar{a}_t^\Lambda - c_t^\Lambda\| \leq K_T^\Lambda \leq \varepsilon. \quad (26)$$

We will now show that  $S_T^w$  is close to an appropriate multiple of  $\|\bar{a}_r^\Lambda - c_r^\Lambda\|$ , for each  $r$  in  $R$ . For  $t$  in  $V$  we have  $w(c_t) \leq w(c_s)$ , and so  $0 \leq w(c_t) - w^r(c_t) \leq w(c_s) - w(c_r) \leq \varepsilon_1/2$  (recall (25)), which gives

$$\left\| \sum_{t \in V} (w(c_t) - w^r(c_t)) (a_t - c_t) \right\| \leq |V| \cdot \frac{\varepsilon_1}{2} \cdot \sqrt{m} \leq \frac{1}{2} \varepsilon_1 \sqrt{m} T.$$

For  $t \notin V$  we have  $0 \leq w(c_t) - w^r(c_t) \leq 1$ , and so (recall (24))

$$\left\| \sum_{t \notin V} (w(c_t) - w^r(c_t)) (a_t - c_t) \right\| \leq (T - |V|) \cdot 1 \cdot \sqrt{m} \leq \varepsilon_2 M \sqrt{m} T.$$

Adding the two inequalities and dividing by  $T$  yields

$$\begin{aligned} S_T^w &= \left\| \frac{1}{T} \sum_{t=1}^T w(c_t) (a_t - c_t) \right\| \leq \left\| \frac{1}{T} \sum_{t=1}^T w^r(c_t) (a_t - c_t) \right\| + \sqrt{m} \left( \frac{\varepsilon_1}{2} + \varepsilon_2 M \right) \\ &\leq \|\bar{a}_r^\Lambda - c_r^\Lambda\| + \sqrt{m} \left( \frac{\varepsilon_1}{2} + \varepsilon_2 M \right), \end{aligned}$$

because  $\sum_{t \leq T} w^r(c_t) (a_t - c_t) = (\sum_{t \leq T} w^r(c_t)) (\bar{a}_r^\Lambda - c_r^\Lambda)$  and  $\sum_{t \leq T} w^r(c_t) \leq T$ . The set  $R$  contains at least  $\varepsilon_2 T$  points (these are all the points in the same cube as  $c_s$ ), i.e.,  $|R| \geq \varepsilon_2 T$ ; averaging over all  $r$  in the set  $R$  and then recalling

(26) finally gives

$$\begin{aligned}
S_T^w &\leq \frac{1}{|R|} \sum_{r \in R} \|\bar{a}_r^\Lambda - c_r^\Lambda\| + \sqrt{m} \left( \frac{\varepsilon_1}{2} + \varepsilon_2 M \right) \\
&\leq \frac{1}{\varepsilon_2 T} \sum_{r \in R} \|\bar{a}_r^\Lambda - c_r^\Lambda\| + \sqrt{m} \left( \frac{\varepsilon_1}{2} + \varepsilon_2 M \right) \\
&\leq \frac{1}{\varepsilon_2} \varepsilon + \sqrt{m} \left( \frac{\varepsilon_1}{2} + \varepsilon_2 M \right) = \frac{\varepsilon \sqrt{L^m}}{\sqrt{\varepsilon}} + \varepsilon_1 \left( \frac{\sqrt{m}}{2} + m^{(m+1)/2} \right) \\
&= \varepsilon_1 \left( 1 + \frac{\sqrt{m}}{2} + m^{(m+1)/2} \right) < K \varepsilon_1 = \varepsilon',
\end{aligned}$$

completing the proof.  $\square$

## 6 Nash Equilibrium Dynamics

In this section we use our results on smooth calibration to obtain dynamics in  $n$ -person games that are in the long run close to Nash equilibria most of the time.

A (finite) *game* is given by a finite set of players  $N$ , and, for each player  $i \in N$ , a finite set of actions<sup>26</sup>  $A^i$  and a payoff function  $u^i : A \rightarrow \mathbb{R}$ , where  $A := \prod_{i \in N} A^i$  denotes the set of action combinations of all players. Let  $n := |N|$  be the number of players,  $m^i := |A^i|$  the number of pure actions of player  $i$ , and  $m := \sum_{i \in N} m^i$ ; also, let  $U$  be a bound on payoffs, i.e.,  $|u^i(a)| \leq U$  for all  $a \in A$  and  $i \in N$ . The set of mixed actions of player  $i$  is  $X^i := \Delta(A^i)$ , the unit simplex (i.e., the set of probability distributions) on  $A^i$ ; we identify the pure actions in  $A^i$  with the unit vectors of  $X^i$ , and so  $A^i \subseteq X^i$ . Put  $C \equiv X := \prod_{i \in N} X^i$  for the set of mixed-action combinations. The payoff functions  $u^i$  are multilinearly extended to  $X$ , and thus  $u^i : X \rightarrow \mathbb{R}$ .

For each player  $i$ , a combination of mixed actions of the other players  $x^{-i} = (x^j)_{j \neq i} \in \prod_{j \neq i} X^j =: X^{-i}$ , and  $\varepsilon \geq 0$ , let  $\text{BR}_\varepsilon^i(x^{-i}) := \{x^i \in X^i : u^i(x^i, x^{-i}) \geq \max_{y^i \in X^i} u^i(y^i, x^{-i}) - \varepsilon\}$  denote the set of  $\varepsilon$ -best replies of  $i$  to  $x^{-i}$ . A (mixed) action combination  $x \in X$  is a *Nash  $\varepsilon$ -equilibrium* if

<sup>26</sup>We refer to one-shot choices as “actions” rather than “strategies,” the latter term being reserved for repeated interactions.

$x^i \in \text{BR}_\varepsilon^i(x^{-i})$  for every  $i \in N$ ; let  $\text{NE}(\varepsilon) \subseteq X$  denote the set of Nash  $\varepsilon$ -equilibria of the game.

A (discrete-time) *dynamic* consists of each player  $i \in N$  playing a pure action  $a_t^i \in A^i$  at each time period  $t = 1, 2, \dots$ ; put  $a_t = (a_t^i)_{i \in N} \in A$ . There is perfect monitoring: at the end of period  $t$  all players observe  $a_t$ . The dynamic is *uncoupled* (Hart and Mas-Colell 2003, 2006, 2013) if the play of every player  $i$  may depend only on player  $i$ 's payoff function  $u^i$  (and not on the other players' payoff functions). Formally, such a dynamic is given by a mapping for each player  $i$  from the history  $h_{t-1} = (a_1, \dots, a_{t-1})$  and his own payoff function  $u^i$  into  $X^i = \Delta(A^i)$  (player  $i$ 's choice may be random); we will call such mappings *uncoupled*. Let  $x_t^i \in X^i$  denote the mixed action that player  $i$  plays at time  $t$ , and put  $x_t = (x_t^i)_{i \in N} \in X$ .

The dynamics we consider are smooth variants of the ‘‘calibrated learning’’ introduced by Foster and Vohra (1997). *Calibrated learning* consists of each player best-replying to calibrated forecasts on the other players' actions; it results in the joint distribution of play converging in the long run to the set of correlated equilibria of the game. Kakade and Foster (2004) defined *publicly calibrated learning*, where each player approximately best-plies to a public weakly calibrated forecast on the joint actions of all players, and proved that most of the time the play is an approximate Nash equilibrium. We consider instead *smooth calibrated learning*, where weak calibration is replaced with the more natural smooth calibration; it amounts to taking calibrated learning and smoothing out both the forecasts and the best replies. Moreover, our forecasts are  $n$ -tuples of mixed strategies (in  $\prod_i \Delta(A^i)$ ), rather than correlated mixtures (in  $\Delta(\prod_i A^i)$ ).

Formally, a *smooth calibrated learning* dynamic is given by:

- (D1) An  $(\varepsilon_c, L_c)$ -smoothly calibrated deterministic procedure, which yields at time  $t$  a forecast  $c_t \in X$  on the distribution of actions of each player.
- (D2) For each player  $i \in N$  an  $L_g$ -Lipschitz  $\varepsilon_g$ -approximate best-reply mapping  $g^i : X \rightarrow X^i$ ; i.e.,  $g^i(x) \in \text{BR}_{\varepsilon_g}^i(x^{-i})$  for every  $x^{-i} \in X^{-i}$ .
- (D3) Each player runs the procedure in (D1), generating at time  $t$  a forecast

$c_t \in X$ ; then each player  $i$  plays at period  $t$  the mixed action<sup>27</sup>  $x_t^i := g^i(c_t) \in X^i$ , where  $g^i$  is given by (D2). All players observe the action combination  $a_t = (a_t^i)_{i \in N} \in A$  that has actually been played, and remember it.

The existence of a deterministic smoothly calibrated procedure in (D1) is given by Theorem 1. For each player  $i$ , the payoff function  $u^i$  is linear in  $x^i$ , and  $|u^i(x^i, x^{-i}) - u^i(y^i, x^{-i})| \leq \sqrt{m^i}U||x^i - y^i||$ , and so  $\mathcal{L}(u^i) \leq \sqrt{m^i}U \leq \sqrt{m}U$ ; the existence of Lipschitz approximate best-reply mappings in (D2) is then given by Lemma 19 in the Appendix (in particular, for  $\varepsilon_g$  and  $L_g$  such that  $L_g \geq \nu_m(\sqrt{m}U/\varepsilon_g)^{m+1}$ ).

Since for each player  $i$  the approximate best reply condition in (D2) makes use *only* of player  $i$ 's payoff function  $u^i$ , we can without loss of generality choose  $g^i$  so as to depend only on  $u^i$ , which makes the dynamic *uncoupled* (see above).

Our result is:

**Theorem 15** *Fix the finite set of players  $N$ , the finite action spaces  $A^i$  for all  $i \in N$ , and the payoff bound  $U < \infty$ . For every  $\varepsilon > 0$ , any smooth calibrated learning dynamic with appropriate parameters<sup>28</sup> is an uncoupled dynamic that satisfies*

$$\liminf_{T \rightarrow \infty} \frac{1}{T} |\{t \leq T : x_t \in \text{NE}(\varepsilon)\}| \geq 1 - \varepsilon \quad (\text{a.s.})$$

for every finite game with payoff functions  $(u^i)_{i \in N}$  that are bounded by  $U$  (i.e.,  $|u^i(a)| \leq U$  for all  $i \in N$  and  $a \in A$ ).

The idea of the proof is as follows. First, assume that the forecasts  $c_t$  are in fact calibrated (rather than just smoothly calibrated) and, moreover, that they are calibrated with respect to the mixed plays  $x_t$  (rather than with respect to the actual plays  $a_t$ ). Because  $x_t$  is given by a fixed function of  $c_t$ ,

<sup>27</sup>Thus  $\mathbb{P}[a_t = a \mid h_{t-1}] = \prod_{i \in N} x_t^i(a^i)$  for every  $a = (a^i)_{i \in N} \in A$ , where  $h_{t-1}$  is the history and  $x_t^i(a^i)$  is the probability that  $x_t^i \in \Delta(A^i)$  assigns to the pure action  $a^i \in A^i$ .

<sup>28</sup>Such as those given in (33).

namely,  $x_t = g(c_t) \equiv (g^i(c_t))_{i \in N}$ , the sequence of mixed plays in those periods when the forecast was a certain  $c$  is the constant sequence  $g(c), \dots, g(c)$ , whose average is  $g(c)$ , and calibration then implies that  $g(c)$  must be close to  $c$  (most of the time, i.e., for forecasts that appear with positive frequency). But we have only smooth calibration; however, because  $g$  is a continuous function, if  $c$  and  $g(c)$  are far from one another then so are  $c'$  and  $g(c')$  for any  $c'$  close to  $c$ , and so the average of such  $g(c')$  is also far from  $c$ , contradicting smooth calibration. Thus, most of the time  $g(c_t)$  is close to  $c_t$ , and hence  $g(g(c_t))$  is close to  $g(c_t)$  (because  $g$  is continuous)—which says that  $g(c_t)$  is close to an approximate best reply to itself, i.e.,  $g(c_t)$  is an approximate Nash equilibrium. Finally, an appropriate use of a strong law of large numbers shows that if the actual plays  $a_t$  are (smoothly) calibrated then so are their expectations, i.e., the mixed plays  $x_t$ . Two crucial features of our dynamic—which are needed to get Nash equilibria, and cannot be obtained with standard, probabilistic, calibration—are, first, that all players always have the same forecast, and second, that (smooth) calibration is preserved despite the fact that the actions depend on the forecasts (leakiness).

**Proof.** This proof goes along similar lines to the proof of Kakade and Foster (2004) for publicly calibrated dynamics (which is the only other calibration-based Nash dynamic to date<sup>29</sup>).

Recall that  $m^i := |A^i|$  and  $m := \sum_{i \in N} m^i$ , and so  $X \subset [0, 1]^m$ . Put  $g(c) := (g^i(c))_{i \in N}$  for every  $c \in X$ ; thus  $g : X \rightarrow X$  is a Lipschitz function with  $\mathcal{L}(g) \leq nL_g$  (because  $\mathcal{L}(g^i) \leq L_g$  for each  $i$ ).

Take  $\Lambda$  to be the  $L_c$ -tent smoothing function:  $\Lambda(c', c) = [1 - L_c||c' - c||]_+$  for all  $c, c' \in X$ .

For each period  $t$ , let  $c_t \in X$  be the *forecast*,  $x_t = g(c_t) \in X$  the *behavior* (i.e., mixed actions), and  $a_t \in A$  the realized pure *actions* ( $c_t, x_t$ , and  $a_t$  all depend on the history). We divide the proof into the following steps: (i) smoothed average actions  $\bar{a}_t^\Lambda$  and forecasts  $\bar{c}_t^\Lambda$  are close (by smooth calibration); (ii) smoothed average actions  $\bar{a}_t^\Lambda$  and behaviors  $\bar{x}_t^\Lambda$  are close (by the law of large numbers); (iii) forecasts  $c_t$  and behaviors  $x_t$  are close (because

---

<sup>29</sup>Recall footnote 4.

smoothing had little effect there); (iv) behaviors  $x_t$  are close to Nash equilibria. Finally, (v) shows how to tweak the parameters to get the desired result.

(i) *Smoothed average actions and smoothed forecasts are close.*

Let  $T_0$  be such that the smooth calibration score  $K_T^\Lambda \leq \varepsilon_c$  for all  $T > T_0$ , i.e.,

$$\frac{1}{T} \sum_{t=1}^T \|\bar{a}_t^\Lambda - c_t^\Lambda\| \leq \varepsilon_c \quad (27)$$

for all  $T > T_0$ .

(ii) *Smoothed average actions and smoothed average behaviors are close.*

Let  $D \subset X$  be a finite  $\varepsilon_1$ -grid of  $X$ . For each  $d \in D$  we have  $\mathbb{E}[\Lambda(c_s, d)a_s \mid h_{s-1}] = \Lambda(c_s, d)x_s$  (given  $h_{s-1}$ , only  $a_s$  is random, and its conditional expectation is  $\mathbb{E}[a_s \mid h_{s-1}] = g(c_s) = x_s$ ). By the Strong Law of Large Numbers for Dependent Random Variables (see Loève 1978, Theorem 32.1.E:  $(1/T) \sum_{s=1}^T (X_s - \mathbb{E}[X_s \mid h_{s-1}]) \rightarrow 0$  as  $T \rightarrow \infty$  a.s., for random variables  $X_s$  that are, in particular, uniformly bounded; note that there are finitely many  $d \in D$ ) we get

$$\lim_{T \rightarrow \infty} \frac{1}{T} \sum_{s=1}^T \Lambda(c_s, d)(a_s - x_s) = 0 \quad \text{for all } d \in D \quad (\text{a.s.}). \quad (28)$$

Thus, for each one of the (almost all) infinite histories  $h_\infty$  where (28) holds, there is a finite  $T_1 \equiv T_1(h_\infty)$  such that  $(1/T) \left\| \sum_{t=1}^T \Lambda(c_s, d)(a_s - x_s) \right\| \leq \varepsilon_1$  for all  $T > T_1$  and all  $d \in D$ . Now for every  $c \in X$  there is  $d \in D$  with  $\|d - c\| \leq \varepsilon_1$ , and so  $|\Lambda(c_s, d) - \Lambda(c_s, c)| \leq L_c \|c - d\| \leq L_c \varepsilon_1$ ; together with<sup>30</sup>  $\|a_s - x_s\| \leq \sqrt{m}$  it follows that

$$\frac{1}{T} \left\| \sum_{s=1}^T \Lambda(c_s, c)(a_s - x_s) \right\| \leq (1 + \sqrt{m}L_c)\varepsilon_1 \quad \text{for all } T > T_1 \text{ and all } c \in X.$$

Taking in particular  $c = c_t$  for all  $t \leq T$ , and then applying Lemma 11 to the

---

<sup>30</sup>Because  $a_s, x_s \in [0, 1]^m$ .

set  $[0, 1]^m$  and  $b_s = a_s - x_s$ , yields

$$\frac{1}{T} \sum_{t=1}^T \|\bar{a}_t^\Lambda - \bar{x}_t^\Lambda\| \leq \gamma_m L_c^{m/2} \sqrt{(1 + \sqrt{m} L_c) \varepsilon_1} =: \varepsilon_2, \quad (29)$$

where the constant  $\gamma_m$  depends only on  $m$ .

(iii) *Behaviors and forecasts are close.*

Because  $\Lambda(c_s, c_t) > 0$  only when  $\|c_s - c_t\| < 1/L_c$ , it follows that  $c_t^\Lambda$ , as a weighted average of such  $c_s$ , satisfies  $\|c_t^\Lambda - c_t\| < 1/L_c$ . Moreover,  $\|x_s - x_t\| = \|g(c_s) - g(c_t)\| \leq nL_g/L_c$ , and so  $\|\bar{x}_t^\Lambda - x_t\| \leq nL_g/L_c$ , which together with (27) and (29) gives

$$\frac{1}{T} \sum_{t=1}^T \|x_t - c_t\| \leq \varepsilon_c + \varepsilon_2 + \frac{1}{L_c} + \frac{nL_g}{L_c} =: \varepsilon_3 \quad (30)$$

for almost every infinite history  $h_\infty$  and for every  $T > \max\{T_0, T_1(h_\infty)\}$ .

(iv) *Behaviors are close to Nash equilibria.* From (30) it immediately follows that, for every  $\varepsilon_4 > 0$ ,

$$\frac{1}{T} |\{t \leq T : \|g(c_t) - c_t\| > \varepsilon_4\}| \leq \frac{1}{\varepsilon_4} \frac{1}{T} \sum_{t=1}^T \|g(c_t) - c_t\| \leq \frac{\varepsilon_3}{\varepsilon_4}. \quad (31)$$

If  $\|g(c_t) - c_t\| \leq \varepsilon_4$  then

$$\|g^i(x_t) - x_t^i\| = \|g^i(g(c_t)) - g^i(c_t)\| \leq L_g \varepsilon_4,$$

and so

$$\begin{aligned} u^i(x_t) &\geq u^i(g^i(x_t), x_t^{-i}) - \sqrt{m^i} U \|g^i(x_t) - x_t^i\| \\ &\geq \max_{y^i \in \Delta(A^i)} u^i(y^i, x_t^{-i}) - \varepsilon_g - \sqrt{m^i} U L_g \varepsilon_4 \end{aligned}$$

(for the second inequality we have used  $g^i(x) \in \text{BR}_{\varepsilon_g}^i(x^{-i})$ ). Therefore  $\|g(c_t) - c_t\| \leq \varepsilon_4$  implies that  $x_t \in \text{NE}(\varepsilon_5)$ , where

$$\varepsilon_5 := \varepsilon_g + \sqrt{m^i} U L_g \varepsilon_4 \quad (32)$$

(recall that  $m = \sum_i m^i \geq m^i$ ), and so, from (30) and (31) we get

$$\frac{1}{T} |\{t \leq T : x_t \notin \text{NE}(\varepsilon_5)\}| \leq \frac{\varepsilon_3}{\varepsilon_4}$$

for all large enough  $T$ , for almost every infinite history.

(v) *Tweaking the parameters.* To bound both  $\varepsilon_3/\varepsilon_4$  and  $\varepsilon_5$  by, say,  $3\varepsilon$ , one may take, for instance (see (29)–(32) and recall Lemma 19 in the Appendix),

$$\begin{aligned} \varepsilon_g &= \varepsilon, \quad L_g = \nu_m \left( \frac{\sqrt{m}U}{\varepsilon} \right)^{m+1}, \\ \varepsilon_4 &= \frac{2\varepsilon}{\sqrt{m}UL_g}, \\ \varepsilon_c &= \varepsilon\varepsilon_4, \quad L_c = \frac{1+nL_g}{\varepsilon\varepsilon_4}, \\ \varepsilon_2 &= \varepsilon\varepsilon_4, \quad \varepsilon_1 = \frac{\varepsilon_2^2}{\gamma_m^2 L_c^m (1 + \sqrt{m}L_c)}, \end{aligned} \tag{33}$$

because we then get  $\varepsilon_5 = \varepsilon + 2\varepsilon = 3\varepsilon$  and  $\varepsilon_3 = \varepsilon\varepsilon_4 + \varepsilon\varepsilon_4 + \varepsilon\varepsilon_4 = 3\varepsilon\varepsilon_4$ .  $\square$

**Remarks.** (a) *Nash dynamics.* Uncoupled dynamics where Nash  $\varepsilon$ -equilibria are played  $1 - \varepsilon$  of the time were first proposed by Foster and Young (2003), followed by Kakade and Foster (2004), Foster and Young (2006), Hart and Mas-Colell (2006), Germano and Lugosi (2007), Young (2009), Babichenko (2012), and others (see also Remark (h) below).

(b) *Coordination.* All players need to coordinate before playing the game on the smoothly calibrated procedure that they will run; thus, at every period  $t$  they all generate the same forecast  $t$ . By contrast, in the original calibrated learning dynamic of Foster and Vohra (1997)—which leads to correlated equilibria—every player may use his own calibrated procedure.

This fits the so-called *Conservation Coordination Law* for game dynamics, which says that some form of “coordination” must be present, either in the limit static equilibrium concept (such as correlated equilibrium) or in the dynamic leading to it (such as Nash equilibrium dynamics). See Hart and Mas-Colell (2003, footnote 19) and Hart (2005, footnote 19).

(c) *Deterministic calibration.* In order for all the players to generate the

same forecasts, it is not enough that they all use the same procedure; in addition, the forecasts must be deterministic (otherwise the randomizations, which are carried out independently by the players, may lead to different actual forecasts). This is the reason that we use smoothly calibrated procedures rather than fully calibrated ones (cf. Oakes 1985 and Foster and Vohra 1998).

(d) *Leaky calibration.* One may use a common *randomized* smoothly calibrated procedure, provided that the randomizations are carried out publicly (i.e., they must be leaked!). Alternatively, a “central bureau of statistics” may be used each period to provide the forecast to all the players.

(e) *Finite memory.* In (D1) one may use a smoothly calibrated procedure that has finite recall and is stationary (see Theorem 1). However, while in the calibration game of Section 5 both the actions  $a_t$  and the forecasts  $c_t$  are monitored and thus become part of the recall window, in the  $n$ -person game only  $a_t$  is monitored (while the forecast  $c_t$  is computed by each player separately, but is *not* played). Therefore, in order to run the calibrated procedure, in the  $n$ -person game each player needs to remember at time  $T$ , in addition to the last  $R$  action combinations  $a_{T-R}, \dots, a_{T-1}$ , also the last  $R$  forecasts  $c_{T-R}, \dots, c_{T-1}$ . “Finite recall” of size  $R$  in the calibration procedure therefore becomes “finite memory” of size  $2R$  in the game dynamic: the memory contains  $R$  elements of  $A$  and  $R$  elements of<sup>31</sup>  $C$ .

Alternatively, to get finite recall rather than finite memory one may introduce an artificial player, say, player 0, with action set  $A^0 := X$  and constant payoff function  $u^0 \equiv 0$ , who plays at each period  $t$  the forecast  $c_t$ , i.e.,  $a_t^0 = c_t$ ; this way the forecasts become part of the recall of all players.

(f) *Forecasting joint play.* In (D1) one may use a procedure that forecasts the joint play: the forecasts  $c_t$  lie in  $\Delta(A)$ , rather than in  $\prod_i \Delta(A^i)$  (the dimension is then larger,  $\prod_i m^i$  instead of  $\sum_i m^i$ ). The approximate best reply functions  $g^i$  can then be defined over  $\Delta(A)$ , and the proof carries through essentially without change. Thus most of the time the play is close to Nash equilibrium, despite the fact that the forecasts are allowed to be

---

<sup>31</sup>For a similar transition from finite recall to finite memory, see Theorem 7 in Hart and Mas-Colell (2006).

correlated; in fact, the forecasts turn out to be close to being independent (because  $g(c_t) \in X$  is independent, and  $c_t$  is close to  $g(c_t)$ ).

(g) *Separate forecasts.* One cannot simplify (D1) by replacing the forecasting procedure that yields  $c_t = (c_t^i)_{i \in N} \in X$  with  $n$  separate forecasting procedures that yield  $c_t^i \in X^i$  for each  $i \in N$ , because then behaviors  $x_t$  and forecasts  $c_t$  need no longer be close (in part (iii) of the proof, when  $c_s^i$  is now close to  $c_t^i$  for some  $i$ , it does not follow that  $c_s^j$  and  $c_t^j$  for  $j \neq i$  are also close, and so neither are  $g(c_s)$  and  $g(c_t)$ ).

(h) *Continuous approximate best reply.* In (D2) one may take the functions  $g^i$  to be continuous rather than Lipschitz and carry the proof with the modulus of continuity instead of the Lipschitz bound (for uncoupledness one would need to require uniform equicontinuity).

(i) *Exhaustive search.* Dynamics that perform exhaustive search can also be used to get the result of Theorem<sup>32</sup> 15. Take for instance a finite grid on  $X$ , say,  $D = \{d_1, \dots, d_M\} \subset X$ , that is fine enough so that there always is a pure Nash  $\varepsilon$ -equilibrium on the grid. Let the dynamic go over the points  $d_1, d_2, \dots$  in sequence until the first time that  $d_T^i \in \text{BR}_\varepsilon^i(d_T^{-i})$  for all  $i$ , following which  $d_T$  is played forever. This is implemented by having for every player  $i$  a distinct action  $a_0^i \in A^i$  that is played at time  $t$  only when  $d_t^i \in \text{BR}_\varepsilon^i(d_t^{-i})$  (otherwise a different action is played); once the action combination  $a_0 = (a_0^i)_{i \in N} \in A$  is played, say, at time  $T$ , each player  $i$  plays  $d_T^i$  at all  $t > T$ . This dynamic is uncoupled (each player only considers  $\text{BR}_\varepsilon^i$ ) and has memory of size 2 (i.e., 2 elements of  $X$ ): for  $t \leq T$ , it consists of  $d_{t-1}$  and  $a_{t-1}$  (the last checked point and the last played action combination); for  $t > T$ , it consists of  $d_T$  and  $a_0$ . Of course, all players need to coordinate before playing the game on the sequence  $d_1, d_2, \dots, d_M$  and the action combination  $a_0$ .

(j) *Continuous action spaces.* The result of Theorem 15 easily extends to continuous action spaces and approximate *pure* Nash equilibria. Assume that for each player  $i \in N$  the set of actions  $A^i$  is a convex compact subset of some Euclidean space (such games arise, for instance, from exchange economies where the actions are net trades; see, e.g., Hart and Mas-Colell 2015). Thus  $A = \prod_{i \in N} A^i$  is a compact convex set in some Euclidean space, say,  $\mathbb{R}^m$ .

---

<sup>32</sup>We thank Yakov Babichenko for suggesting this.

For every  $\varepsilon \geq 0$ , the set of *pure  $\varepsilon$ -best replies* of player  $i$  to  $a^{-i} \in A^{-i}$  is  $\text{PBR}_\varepsilon^i(a^{-i}) := \{a^i \in A^i : u^i(a^i, a^{-i}) \geq \max_{b^i \in A^i} u^i(b^i, a^{-i}) - \varepsilon\}$ . An action combination  $a \in A$  is a *pure Nash  $\varepsilon$ -equilibrium* if  $a^i \in \text{PBR}_\varepsilon^i(a^{-i})$  for every  $i \in N$ ; let  $\text{PNE}(\varepsilon) \subseteq A$  denote the set of pure Nash  $\varepsilon$ -equilibria.

Smooth calibrated learning is defined as above, except that now the approximate best replies are pure actions (the play is  $a_t = g(c_t)$ , and it is monitored by all players). Our result here is:

**Theorem 16** *Fix the finite set of players  $N$ , the convex compact action spaces  $A^i$  for all  $i \in N$ , and the Lipschitz bound  $L < \infty$ . For every  $\varepsilon > 0$ , and any smooth calibrated learning dynamic with appropriate parameters, there is  $T_0 \equiv T_0(\varepsilon, L)$  such that for every  $T \geq T_0$ ,*

$$\frac{1}{T} |\{t \leq T : a_t \in \text{PNE}(\varepsilon)\}| \geq 1 - \varepsilon$$

*for every game with payoff functions  $(u^i)_{i \in N}$  that are  $L$ -Lipschitz (i.e.,  $\mathcal{L}(u^i) \leq L$ ) and quasi-concave in one's own action (i.e.,  $u^i(a^i, a^{-i})$  is quasi-concave in  $a^i \in A^i$  for every  $a^{-i} \in A^{-i}$ ), for all  $i \in N$ .*

**Proof.** We now have  $A^i = X^i$  and  $a_t = x_t = g(c_t)$ , and everything is deterministic. Proceed as in the proof of Theorem 15, skipping part (ii) (the use of the Law of Large Numbers) and taking  $\varepsilon_1 = 0$  and  $T_1 = 0$ .  $\square$

(*k*) *Reaction function and fixed points.* The proof of Theorem 15 shows that in the leaky calibration game, if the A-player uses a stationary strategy given by a Lipschitz “reaction” function  $g$  (i.e., he plays  $g(c_t)$  at time  $t$ ), then smooth calibration implies that the forecasts  $c_t$  are close to fixed points of  $g$  most of the time.

## References

Azoury, K. S. and M. K. Warmuth (2001), “Relative Loss Bounds for On-Line Density Estimation with the Exponential Family of Distributions,” *Machine Learning* 43, 211–246.

- Babichenko, Y. (2012), “Completely Uncoupled Dynamics and Nash Equilibria,” *Games and Economic Behavior* 76, 1–14.
- Cesa-Bianchi, N. and G. Lugosi (2006), *Prediction, Learning, and Games*, Cambridge University Press.
- Dawid, A. P. (1982), “The Well-Calibrated Bayesian,” *Journal of the American Statistical Association* 77, 605–613.
- Dawid, A. P. (1985), “Self-Calibrating Priors Do Not Exist: Comment,” *Journal of the American Statistical Association* 80, 340–341.
- Foster, D. P. (1991), “Prediction in the Worst Case,” *The Annals of Statistics* 19, 1084–1090.
- Foster, D. P. (1999), “A Proof of Calibration via Blackwell’s Approachability Theorem,” *Games and Economic Behavior* 29, 73–78.
- Foster, D. P. and S. M. Kakade (2006), “Calibration via Regression,” *IEEE Information Theory Workshop 2006*.
- Foster, D. P. and R. V. Vohra (1997), “Calibrated Learning and Correlated Equilibrium,” *Games and Economic Behavior* 21, 40–55.
- Foster, D. P. and R. V. Vohra (1998), “Asymptotic Calibration,” *Biometrika* 85, 379–390.
- Foster, D. P. and H. P. Young (2003), “Learning, Hypothesis Testing, and Nash Equilibrium,” *Games and Economic Behavior* 45, 73–96.
- Foster, D. P. and H. P. Young (2006), “Regret Testing: Learning to Play Nash Equilibrium without Knowing You Have an Opponent,” *Theoretical Economics* 1, 341–367.
- Foster, J. (1999), “On Relative Loss Bounds in Generalized Linear Regression,” in *12th International Symposium on Fundamentals of Computation Theory (FCT ’99)*, 269–280.
- Germano, F. and G. Lugosi (2007), “Global Nash Convergence of Foster and Young’s Regret Testing,” *Games and Economic Behavior* 60, 135–154.
- Hart, S. (2005), “Adaptive Heuristics,” *Econometrica* 73, 1401–1430; Chapter 11 of Hart and Mas-Colell (2013).

- Hart, S. and A. Mas-Colell (2000), “A Simple Adaptive Procedure Leading to Correlated Equilibrium,” *Econometrica* 68, 1127–1150; Chapter 2 of Hart and Mas-Colell (2013).
- Hart, S. and A. Mas-Colell (2001), “A General Class of Adaptive Strategies,” *Journal of Economic Theory* 98, 26–54; Chapter 3 of Hart and Mas-Colell (2013).
- Hart, S. and A. Mas-Colell (2003), “Uncoupled Dynamics Do Not Lead to Nash Equilibrium,” *American Economic Review* 93, 1830–1836; Chapter 7 of Hart and Mas-Colell (2013).
- Hart, S. and A. Mas-Colell (2006), “Stochastic Uncoupled Dynamics and Nash Equilibrium,” *Games and Economic Behavior* 57, 286–303; Chapter 8 of Hart and Mas-Colell (2013).
- Hart, S. and A. Mas-Colell (2013), *Simple Adaptive Strategies: From Regret-Matching to Uncoupled Dynamics*, World Scientific, 2013.
- Hart, S. and A. Mas-Colell (2015), “Markets, Correlation, and Regret-Matching,” *Games and Economic Behavior* 93, 42–54.
- Hazan, E. and S. Kakade (2012), “(Weak) Calibration is Computationally Hard,” *25th Annual Conference on Learning Theory (COLT '12); JMLR: Workshop and Conference Proceedings* 23 (2012), 3.1–3.10.
- Kakade, S. M. and D. P. Foster (2004), “Deterministic Calibration and Nash Equilibrium,” in *17th Annual Conference on Learning Theory (COLT '04)*; *Journal of Computer and System Sciences* 74 (2008), 115–130.
- Lehrer, E. and E. Solan (2009), “Approachability with Bounded Memory,” *Games and Economic Behavior* 66, 995–1004.
- Loève, M. (1978), *Probability Theory, Vol. II*, 4th edition, Springer.
- Oakes, D. (1985), “Self-calibrating Priors Do Not Exist,” *Journal of the American Statistical Association* 80, 339.
- Olszewski, W. (2015), “Calibration and Expert Testing,” in *Handbook of Game Theory, Vol. 4*, H. P. Young and S. Zamir (editors), Springer, 949–984.
- Papadimitriou, C. (1994). “On the Complexity of the Parity Argument and Other Inefficient Proofs of Existence,” *Journal of Computer and System Sciences* 48, 498–532.

Vovk, V. (2001), “Competitive On-Line Statistics,” *International Statistical Review* 69, 213–248.

Young, H. P. (2009), “Learning by Trial and Error,” *Games and Economic Behavior* 65, 626–643.

## A Appendix

Let  $C$  be a compact subset of  $\mathbb{R}^m$  and let  $\varepsilon > 0$ . A *maximal*  $2\varepsilon$ -net in  $C$  is a maximal collection of points  $z_1, \dots, z_K \in C$  such that  $\|z_k - z_j\| \geq 2\varepsilon$  for all  $k \neq j$ ; maximality implies  $\cup_{k=1}^K B(z_k, 2\varepsilon) \supseteq C$ . Let  $\alpha_k(x) := [3\varepsilon - \|x - z_k\|]_+$ , and put  $\bar{\alpha}(x) := \sum_{k=1}^K \alpha_k(x)$ . For every  $x \in C$  we have  $0 \leq \alpha_k(x) \leq 3\varepsilon$  and  $\bar{\alpha}(x) \geq \varepsilon$  (since  $\alpha_k(x) \geq \varepsilon$  when  $x \in B(z_k, 2\varepsilon)$ , and the union of these balls covers  $C$ ). Finally, define  $\beta_k(x) := \alpha_k(x)/\bar{\alpha}(x)$ .

**Lemma 17** *The functions  $(\beta_k)_{1 \leq k \leq K}$  satisfy the following properties:*

- (i)  $\beta_k(x) \geq 0$  for all  $x \in C$  and all  $k$ .
- (ii)  $\sum_{k=1}^K \beta_k(x) = 1$  for all  $x \in C$ .
- (iii)  $\beta_k(x) = 0$  for all  $x \notin B(z_k, 3\varepsilon)$ .
- (iv) For each  $x \in C$  there are at most<sup>33</sup>  $4^m$  indices  $k$  such that  $\beta_k(x) > 0$ .
- (v)  $\mathcal{L}(\beta_k) \leq 4^{m+2}/\varepsilon$  for every  $k$ .

**Proof.** (i) and (ii) are immediate. For (iii), we have  $\beta_k(x) > 0$  iff  $\alpha_k(x) > 0$  iff  $\|x - z_k\| < 3\varepsilon$ . This implies that  $B(z_k, \varepsilon) \subseteq B(x, 4\varepsilon)$ . The open balls of radius  $\varepsilon$  with centers at  $z_k$  are disjoint (because  $\|z_k - z_j\| \geq 2\varepsilon$  for  $k \neq j$ ), and so there can be at most  $4^m$  such balls included in  $B(x, 4\varepsilon)$  whose volume

---

<sup>33</sup>We have not tried to get the best bounds in (iv) and (v); indeed, they may be easily reduced.

is  $4^m$  times larger; this proves (iv). For every  $x, y \in C$ :

$$\begin{aligned}
|\beta_k(x) - \beta_k(y)| &\leq \left| \frac{\alpha_k(x)}{\bar{\alpha}(x)} - \frac{\alpha_k(y)}{\bar{\alpha}(x)} \right| + \left| \frac{\alpha_k(y)}{\bar{\alpha}(x)} - \frac{\alpha_k(y)}{\bar{\alpha}(y)} \right| \\
&\leq \frac{1}{\bar{\alpha}(x)} |\alpha_k(x) - \alpha_k(y)| + \frac{\alpha_k(y)}{\bar{\alpha}(x)\bar{\alpha}(y)} \sum_{j=1}^K |\alpha_j(x) - \alpha_j(y)| \\
&\leq \frac{1}{\varepsilon} \|x - y\| + \frac{3\varepsilon}{\varepsilon \cdot \varepsilon} 2 \cdot 4^m \|x - y\| \leq \frac{4^{m+2}}{\varepsilon} \|x - y\|
\end{aligned}$$

(since  $\bar{\alpha}(x) \geq \varepsilon$ ,  $\alpha_k(x) \leq 3\varepsilon$ , and there are at most  $2 \cdot 4^m$  indices  $j$  where neither  $\alpha_j(x)$  nor  $\alpha_j(y)$  vanishes); this proves (v).  $\square$

Thus, the functions  $(\beta_k)_{1 \leq k \leq K}$  constitute a *Lipschitz partition of unity* that is subordinate to the maximal  $2\varepsilon$ -net  $z_1, \dots, z_K$ . Next, we obtain a basis for the Lipschitz functions on  $C$ .

**Lemma 18** *Let  $W_L$  be the set of functions  $w : C \rightarrow [0, 1]$  with  $\mathcal{L}(w) \leq L$ . Then for every  $\varepsilon > 0$  there exist  $d$  functions  $f_1, \dots, f_d \in W_L$  such that for every  $w \in W_L$  there is a vector  $\varpi \equiv \varpi_w \in [0, 1]^d$  satisfying*

$$\max_{x \in C} \left| w(x) - \sum_{i=1}^d \varpi_i f_i(x) \right| < \varepsilon.$$

Moreover, one can take  $d = O(L^m/\varepsilon^{m+1})$ .

**Proof.** Put  $\varepsilon_1 := \varepsilon/(3L)$ . Let  $z_1, \dots, z_K$  be a maximal  $2\varepsilon_1$ -net on  $C$ , and let  $\beta_1, \dots, \beta_K$  be the corresponding Lipschitz partition of unity given by Lemma 17 (for  $\varepsilon_1$ ).

Given  $w \in W_L$ , let  $v(x) := \sum_{k=1}^N w(z_k)\beta_k(x)$ ; then  $w(z_k) \in [0, 1]$  and we have

$$\begin{aligned}
|w(x) - v(x)| &= \left| \sum_{k=1}^N (w(x) - w(z_k)) \beta_k(x) \right| \leq \sum_{k:\beta_k(x)>0} \beta_k(x) |w(x) - w(z_k)| \\
&\leq \sum_{k:\beta_k(x)>0} \beta_k(x) 3\varepsilon_1 L = 3\varepsilon_1 L,
\end{aligned}$$

since  $\beta_k(x) > 0$  implies  $\|x - z_k\| < 3\varepsilon_1$  and thus  $|w(x) - w(z_k)| \leq L\|x - z_k\| \leq L \cdot 3\varepsilon_1$  (because  $\mathcal{L}(w) \leq L$ ).

Now  $\mathcal{L}(\beta_k) \leq 4^{m+2}/\varepsilon_1$  by (v) of Lemma 17; we thus replace each  $\beta_k$  by the sum of  $Q = \lceil 4^{m+2}/(\varepsilon_1 L) \rceil$  identical copies of  $(1/Q)\beta_k$ —denote them  $f_{k,1}, \dots, f_{k,Q}$ —which thus satisfy  $\mathcal{L}(f_{k,q}) = (1/Q)\mathcal{L}(\beta_k) \leq L$ , and so

$$\left| w(x) - \sum_{k=1}^K \sum_{q=1}^Q w(z_k) f_{k,q}(x) \right| = |w(x) - v(x)| \leq 3\varepsilon_1 L = \varepsilon.$$

The  $d = KQ$  functions  $(f_{k,q})_{1 \leq k \leq K, 1 \leq q \leq Q}$  yield our result.

Finally,  $K = O(\varepsilon_1^{-m})$  (because  $C$  contains the  $K$  disjoint open balls of radius  $\varepsilon_1$  centered at the  $z_k$ ) and  $Q \leq 4^{m+2}/(\varepsilon_1 L) + 1$ , and so  $d = KQ = O(\varepsilon_1^{-m-1} L^{-1}) = O(\varepsilon^{-m-1} L^m)$ .  $\square$

In the game setup we construct  $\varepsilon$ -best reply functions that are Lipschitz. The following lemma applies when the action spaces are finite (as in Theorem 15), and also when they are continuous (as in Theorem 16). In the former  $C = X = \prod_{i \in N} X^i$  where  $X^i = \Delta(A^i)$ , and in the latter  $C = X = A = \prod_{i \in N} A^i$ , and the set  $\Delta(A^i)$  is identified with  $A^i$ ; also,  $\text{BR}_\varepsilon^i$  stands for  $\text{PBR}_\varepsilon^i$ , the set of *pure*  $\varepsilon$ -best replies.

**Lemma 19** *Assume that for each player  $i \in N$  the function  $u^i : X \rightarrow \mathbb{R}$  is a Lipschitz function with  $\mathcal{L}(u^i) \leq L$ , and  $u^i(\cdot, c^{-i})$  is quasi-concave on  $X^i$  for every fixed  $c^{-i} \in X^{-i}$ . Then for every  $\varepsilon > 0$  there is a Lipschitz function  $g^i : X \rightarrow X^i$  such that  $g^i(c) \in \text{BR}_\varepsilon^i(c^{-i})$  for all  $c \in X$ , and  $\mathcal{L}(g^i) \leq \nu_m(L/\varepsilon)^{m+1}$  where the constant  $\nu_m$  depends only on the dimension  $m$ .*

**Proof.** Put  $\varepsilon_1 := \varepsilon/(6L)$ . Let  $z_1, \dots, z_K \in C$  be a maximal  $2\varepsilon_1$ -net on  $C$ , and let  $\beta_1, \dots, \beta_K$  be the subordinated Lipschitz partition of unity given by Lemma 17. For each  $i \in N$  and  $1 \leq k \leq K$  take  $x_k^i \in \text{BR}_0^i(z_k^{-i})$ , and define  $g^i(c) := \sum_{k=1}^K \beta_k(c) x_k^i$ . Because  $\beta_k(c) > 0$  if and only if  $\|c - z_k\| < 3\varepsilon_1$ , it follows that  $x_k^i \in \text{BR}_\varepsilon^i(c^{-i})$  (indeed, for every  $y^i \in \Delta(A^i)$  we have  $u^i(x_k^i, c^{-i}) > u^i(x_k^i, z_k^{-i}) - 3L\varepsilon_1 \geq u^i(y^i, z_k^{-i}) - 3L\varepsilon_1 > u^i(y^i, c^{-i}) - 6L\varepsilon_1 = \varepsilon$ , where we have used  $\mathcal{L}(u^i) \leq L$  twice, and  $x_k^i \in \text{BR}_0^i(z_k^{-i})$ ). The set  $\text{BR}_\varepsilon^i(c^{-i})$

is convex by the quasi-concavity assumption, and so  $g^i(c)$ , as an average of such  $x_k^i$ , belongs to  $\text{BR}_\varepsilon^i(c^{-i})$ .

Now  $\max_{c \in C} \|c\| \leq \sqrt{m}$  (because  $C \subseteq [0, 1]^m$ ), and so  $\|x_k\| \leq \sqrt{m}$  (where  $x_k = (x_k^i)_{i \in N}$ ) for all  $k$ , and  $K \leq (\sqrt{m}/\varepsilon_1)^m$  (because  $C \subseteq B(0, \sqrt{m})$  contains the  $K$  disjoint open balls of radius  $\varepsilon_1$  centered at the points  $z_k$ ). Therefore the Lipschitz constant of  $g(c) = \sum_{k=1}^K \beta_k(c)x_k$  satisfies, by Lemma 17 (v),  $\mathcal{L}(g) \leq \sum_{k=1}^K \|x_k\| \mathcal{L}(\beta_k) \leq (\sqrt{m}/\varepsilon_1)^m \sqrt{m} 4^{m+2}/\varepsilon_1 = \nu_m \varepsilon^{-m-1} L^{m+1}$  for  $\nu_m = \sqrt{m}^{m+1} 4^{m+2} 6^{m+1}$ .  $\square$